# 8,2016

## MODEL CALCULATIONS OF SHORT-TERM FORECASTS OF RUSSIAN ECONOMIC TIME SERIES

M.Turuntseva, E.Astafieva, M.Bayeva, A.Bozhechkova, A.Buzaev, T.Kiblitskaya, Yu.Ponomarev and A.Skrobotov

INTRODUCTION TO ALL THE ISSUES	2
INDUSTRIAL PRODUCTION AND RETAIL SALES	5
FOREIGN TRADE INDICES	7
DYNAMICS OF PRICES	7
MONETARY INDICES	10
International reserves	11
FOREIGN EXCHANGE RATES	12
THE LIVING STANDARD INDICES	12
EMPLOYMENT AND UNEMPLOYMENT	13
ANNFX	14

### INTRODUCTION TO ALL THE ISSUES

This paper presents calculations of various economic indicators for the Russian Federation in the period from September 2016 to February 2017, which were performed using time series models developed as a result of research conducted by the Gaidar Institute over the past few years¹. A method of forecasting falls within the group of formal or statistical methods. In other words, the calculated values neither express the opinion nor expert evaluation of the researcher, rather they are calculations of future values for a specific economic indicator, which were performed using formal ARIMA models (p, d, q) given a prevailing trend and its, in some cases, significant changes. The presented forecasts are of inertial nature, because respective models rely upon the dynamics of the data registered prior to the moment of forecasting and depend too heavily on the trends, which are typical of the time series in the period immediately preceding the time horizon to be forecast. The foregoing calculations of future values of economic indicators for the Russian Federation can be used in making decisions on economic policy, provided that the general trends, which were seen prior to forecasting for each specific indicator, remain the same, i.e. prevailing long-term trends will see no serious shocks or changes in the future.

Despite that there is a great deal of data available on the period preceding the crisis of 1998, models of forecasting were analyzed and constructed using only the time horizon which followed August 1998. This can be explained by the findings of previous studies<sup>2</sup> which concluded, among other key inferences, that the quality of forecasts was deteriorated in most of the cases when the data on the pre-crisis period was used. Additionally, it currently seems incorrect to use even shorter series (following the crisis of 2008), because statistical characteristics of models based on such a short time horizon are very poor.

Models for the economic indicators in question were evaluated using standard methods of time series analysis. Initially, the correlograms of the studied series and their first differences were analyzed in order to determine the maximum number of delayed values to be included into the specifications of a model. Then, the results of analyzed correlograms served as the basis for testing all the series for weak stationarity (or stationarity around the trend) using the Dickey–Fuller test. In some cases, the series were tested for stationarity around the segmented trend using Perron and Zivot–Andrews tests for endogenous structural changes<sup>3</sup>.

The series were broken down into weak stationary, stationary near the trend, stationary near the trend with structural change or difference stationary, and then models, which corresponded to each type (regarding the levels and including, if necessary, the trend or segmented trend or differences), were evaluated. The Akaike and Schwartz information criteria, the properties of models' residuals (lack of autocorrelation, homoscedasticity and normality) and the quality of the in-sample-forecasts based on these models were used to choose the best model. Forecast values were calculated for the best of the models constructed for each economic indicator.

Additionally, the Bulletin presents future monthly values of the CPI, which were calculated using models developed at the Gaidar Institute, and volumes of imports/exports from/to all countries, which were calculated using structural models (SM). The forecast values based on the structural models may, in some cases, produce better results than ARIMA-models do, because structural

<sup>1</sup> See, for example, R.M. Entov, S.M. Drobyshevsky, V.P. Nosko, A.D. Yudin. The Econometric Analysis of the Time Series of the Main Macroeconomic Indices. Moscow, IET, 2001; R.M. Entov, V.P. Nosko, A.D. Yudin, P.A. Kadochnikov, S.S. Ponomarenko. Problems of Forecasting of Some Macroeconomic Indices. Moscow, IET, 2002; V. Nosko, A. Buzaev, P. Kadochnikov, S. Ponomarenko. Analysis of the Forecasting Parameters of Structural Models and Models with the Outputs of the Polls of Industries. Moscow, IET, 2003; M.Yu. Turuntseva and T.R. Kiblitskaya, Qualitative Properties of Different Approaches to Forecasting of Social and Economic Indices of the Russian Federation. Moscow, IET, 2010.

<sup>2</sup> Ibid.

<sup>3</sup> See.: Perron, P. Further Evidence on Breaking Trend Functions in Macroeconomic Variables, *Journal of Econometrics*, 1997, 80, pp. 355–385; Zivot, E. and D.W.K. Andrews. Further Evidence on the Great Crash, the Oil-Price Shock, and Unit-Root Hypothesis. *Journal of Business and Economic Statistics*, 1992, 10, pp. 251–270.

### INTRODUCTION TO ALL THE ISSUES

models are constructed by adding information of the dynamics of exogenous variables. Besides, the use of structural forecasts in making aggregated forecasts (i.e. forecasts obtained as average value from several models) may help make forecast values more accurate.

The dynamics of the Consumer Price Index was modeled using theoretical assumptions arising from the monetary theory. The following was used as explanatory variables: money supply, output volume, the dynamics of the ruble-dollar exchange rate, which reflects the dynamics of alternative cost of money-keeping. The model for the Consumer Price Index also included the price index in the electric power industry, because the dynamics of manufacturers' costs relies heavily on this indicator.

The baseline indicator to be noted is the real exchange rate, which can influence the value of exports and imports, and its fluctuations can result in changes to the relative value of domestically-produced and imported goods, though the influence of this indicator turns out to be insignificant in econometric models. Global prices of exported resources, particularly crude oil prices, are most significant factors, which determine the dynamics of exports: a higher price leads to greater exports of goods. The level of personal income in the economy (labor costs) was used to describe the relative competitive power of Russian goods. Fictitious variables D12 and D01 – equal to one in December and January and zero in other periods – were added so that seasonal fluctuations were factored in. The dynamics of imports is effected by personal and corporate incomes whose increase triggers higher demand for all goods including imported ones. The real disposable cash income reflects the personal income; the Industrial Production Index reflects the corporate income.

The forecast values of foreign exchange rates were also calculated using structural models of their dependence on global crude oil prices.

The forecast values of explanatory variables, which are required for forecasting on the basis of structural models, were calculated using ARIMA-models (p, d, q).

The paper also presents calculations of the values of the Industrial Production Index, the Producer Price Index and the Total Unemployment Index, which were calculated using the results of business surveys conducted by the Gaidar Institute. Empirical studies show¹ that the use of series of business surveys as explanatory variables ² in forecasting models can make forecasting more accurate on the average. Future values of these indicators were calculated using ADL-models (seasonal autoregressive delays were added).

The Consumer Price Index and the Producer Price Index are also forecast using large datasets (factor models – FM). The construction of factor models relies basically on the evaluation of the principal components of a large dataset of socio-economic indicators (112 indicators in this case). The lags of these principal components and the lags of the explanatory variable are used as explanatory variables in these models. A quality analysis of the forecasts obtained for different configurations of the factor models was used to chose a model for the CPI, which included 9<sup>th</sup>, 12<sup>th</sup> and 13<sup>th</sup> lags of the four principal components, as well as 1<sup>st</sup> and 12<sup>th</sup> lags of the variable itself, and a model for the PPI, which included 8<sup>th</sup>, 9<sup>th</sup> and 12<sup>th</sup> lags of the four principal components, as well as 1st, 3<sup>rd</sup> and 12<sup>th</sup> lags of the variable itself.

All calculations were performed using the Eviews econometric package.

<sup>1</sup> See, for example: V. Nosko, A. Buzaev, P. Kadochnikov, S. Ponomarenko. *The Analysis of Forecasting Parameters of Structural Models and Models with Business Surveys' Findings*. Moscow, IEP, 2003.

Used as explanatory variables were the following series of the business surveys: the current/expected change in production, the expected changes in the solvent demand, the current/expected price changes and the expected change in employment.

CALCULATIONS OF FORECAST VALUES OF INDICES OF INDUSTRIAL PRODUCTION, 1 (%)

IIP	for machinery	<b>E</b>	NBU HS		-2.2	-4.9	-2.7	2.4	14.0	15.1		-1.7	-7.3	9.5	-11.7	-10.3	-9.5
	for ma	7	Rossta		0.9	-7.8	-3.4	-12.2	16.6	18.8		-9.7	-5.9	-5.8	-10.9	2.2	10
P mary	s and d metal ucts	Э.Е	NBU HS		-4.4	8.0-	1.3	1.9	8.9	3.0		-1.6	-4.0	-4.0	-6.5	-7.4	0.3
IIP for primary	fabricated meta	1	Rossta		-2.5	0.4	4.8	8.9	10.9	8.8		-3.7	-6.9	-6.5	-13.3	-6.5	-1.4
Д	e and leum	Э.Е	NBU HS		-1.8	1.9	9.0-	-3.1	0.4	-1.5		-3.0	-5.8	-3.8	9.0-	-6.8	-4.1
III	ior coke and petroleum	7	Rossta		-2.5	-0.2	-2.5	-4.5	-1.0	-1.7	14/2015	-0.5	-5.1	-1.8	1.9	-3.2	-1.9
Ъ	products	3E	NBU HS	ous year	1.4	1.7	2.0	3.1	3.2	1.7	onth of 20	0.0	9.0	1.3	1.0	3.0	4.9
П	for food products	7	Rossta	Expected growth on the respective month of the previous year	2.1	1.8	2.2	2.2	3.1	2.7	For reference: actual growth in 2015/2016 on the respective month of 2014/2015	1.4	1.5	2.4	3.1	2.0	4.5
IIP for utilities (elec-	tricity, water, and gas)	Э.Е	NBU HS	month of	2.4	0.0	-0.1	0.9	-1.7	7.1	n the resp	-2.0	-3.1	-2.7	-6.4	4.6	6.0-
II for utilit	tricity, wat and gas)	7	Rossta	espective	5.2	1.2	1.5	4.4	-2.3	1.0	15/2016 o	-2.6	-3.6	-3.5	-6.3	2.5	0.0
IIP	ing	3E	NBU HS	on the re	-1.5	-1.2	0.2	-2.2	0.1	-2.5	vth in 20	-3.5	-5.0	-3.3	-4.2	-4.5	-0.1
	ior man	7	Rossta	ed growth	-0.5	-1.1	-0.9	-2.4	-1.1	-3.1	ctual grov	-5.4	-5.9	-5.3	-6.1	-5.6	-1.0
IIP	for mining	Э.Е	NBU HS	Expect	1.4	1.3	1.9	2.3	1.3	-2.6	erence: a	2.8	2.4	1.7	1.3	8.0	6.4
	for m	7	Rossta		-2.4	-2.4	-0.5	0.0	1.1	-1.2	For ref	8.0	1.4	-0.1	0.1	0.4	×0
	on.	HSE	BR		0.7	0.2	1.6	9.0	2.0	9.0		-1.3	-2.4	.7	-2.9	.1	1.9
lex	of industrial production	NRU HSE	AMIЯA		0.3	0.4	1.4	1.8	9.0	-0.4		-1	-2	-1.7	-2	-1.1	-
Index	industria.	stat	BS		1.0	0.3	1.8	0.7	2.1	6.0		-3.7	-3.6	-3.5	-4.5	-2.7	0
	Jo	Rosstat	AMIAA		1.3	0.0	1.3	9.0-	2.0	1.7		-3	6-	-3	4-	-2	1.0
					Sep 16	Oct 16	Nov 16	Dec 16	Jan 17	Feb 17		Sep 15	Oct 15	Nov 15	Dec 15	Jan 16	Feb 16

Note. In the time spans under review, the series of the Rosstat and the NRU HSE chain indices of IIP, as well as the NRU HSE chain IIP for manufacturing are identified as the NRU HSE chain IIP for mining and Rosstat chain IIP for mining and the second the machinery and equipment are identified as stationary processes around the stationary processes around the trend with an endogenous structural change; the series of the Rosstat and the NRU HSE chain IIPs for manufacturing, for primary metals and trend with two endogenous structural changes. The time series of other chain indices are stationary at levels.

It is to be noted that for making of forecasts so-called "raw" indices (without seasonal and calendar adjustment) were used and for that reason in most models existence of the season factor is taken into account and, as a consequence, the obtained outputs reflect the seasonal dynamics of the series.

### INDUSTRIAL PRODUCTION AND RETAIL SALES

### **Industrial production**

For making forecast for September 2016 – February 2017, the series of monthly data of the indices of industrial production released by the Federal State Statistics Service (Rosstat) from January 2002 to June 2016, as well as the series of the base indices of industrial production released by the National Research University Higher School of Economics (NRU HSE¹) over the period from January 1999 to July 2016 were used (the value of January 2000 was equal to 100%). The forecast values of the series were calculated on the basis of ARIMA- models. The forecast values of the Rosstat and the NRU HSE indices of industrial production are calculated using business surveys (BS) as well. The obtained results are shown in Table 1.

As seen from  $Table\ 1$ , the average<sup>2</sup> growth of the index of industrial production computed by the NRU HSE in September 2016 – February 2017 against the same period of the previous year amounts to 0.8% on industry as a whole. For the index of industrial production computed by Rosstat, this indicator constitutes 1.0%. As of end-2016, the forecast annual growth of the index of industrial production computed by Rosstat will amount to 0.3%, and the index of industrial production computed by the NRU HSE – 0.6%.

The average monthly values of the index of industrial production for mining computed by Rosstat and the NRU HSE in September 2016-February 2017 will come to (-0.9%) and 0.9%, respectively. In production of coke and petroleum products growth is forecast at (-2.1%) and (-0.8%) for Rosstat and the NRU HSE indices, respectively.

In September 2016 – February 2017 in comparison with the same period of last year, the average growth of the NRU HSE index of industrial production in manufacturing comes to (-1.2%) and the Rosstat index at (-1.5%). The average monthly values of the Rosstat and the NRU HSE index for industrial production of food products constitute 2.3% and 2.2%, respectively. The average monthly values

of the index of industrial production for primary metals and fabricated metal products in September 2016 – February 2017 computed by Rosstat and the NRU HSE constitute 5.2% and 1.3%, respectively. In manufacture of machinery and equipment, the average increase is forecast at 3.0% and 3.6% for the Rosstat and the NRU HSE indices, respectively.

The average growth of the index of industrial production for utilities (electricity, gas and water) computed by Rosstat for September 2016 – February 2017 in comparison with the same period of the previous year constitutes 1.8%; the same indicator for the NRU HSE index comes to 2.3%.

### **Retail Sales**

This section (Table 2) presents forecasts of monthly retail sales made on the basis of monthly Rosstat data over January 1999 – August 2016.

As seen from *Table 2*, the average forecast growth of nominal volumes of monthly trade turnover

Table 2
CALCULATIONS OF FORECAST VALUES OF THE
RETAIL SALES AND THE REAL RETAIL SALES

IXL17	AIL SALLS AND THE KLA	IL KLIAIL SALLS
Fo	precast value according to	ARIMA-model
	Retail sales, billion RUR	Real retail sales (as
	(in brackets – growth on	•
	the respective month of	
	the previous year, %)	ous year)
Sep 16	2,387.6 (2.8)	96.2
Oct 16	2,456.8 (3.0)	97.4
Nov 16	2,480.0 (3.9)	97.8
Dec 16	3,128.3 (7.9)	96.9
Jan 17	2,141.6 (0.7)	96.6
Feb 17	2,115.3 (0.8)	95.0
For r	eference: actual values in	the same months
	of 2015/2016	
Sep 15	2,322.6	89.3
Oct 15	2,385.2	88.7
Nov 15	2,387.3	87.8
Dec 15	2,898.1	85.9
Jan 16	2,126.2	93.6
Feb 16	2,098.6	95.3

**Note**. The series of retail sales and real retail sales over January 1999 – August 2016.

<sup>1</sup> The indices in question are calculated by E.A. Baranov and V.A. Bessonov.

<sup>2</sup> The average growth of industrial production indices is understood here as the average value of the said indices for six forecast months.

Table 3

CALCULATIONS OF FORECAST VALUES OF VOLUMES OF FOREIGN TRADE TURNOVER WITH COUNTRIES OUTSIDE THE CIS

		() ()			77.72								5 5 5	L I I ()		
	H	export to a	Export to all countries	3	Im	port from	Import from all countries	Se	Export t	o countrie	Export to countries outside the CIS	ne CIS	Import fr	Import from countries outside the CIS	es outside	the CIS
	Forecast values (billion USD a month)	t values USD a ith)	Percentage of actual data in the respective month of the previ- ous year	ntage data in pective the previ-	Forecast values (billion USD a month)	t values USD a nth)	Percentage of actual data in the respective month of the previous year	ntage data in oective he previ-	Forecast values (billion USD a month)	values USD a th)	Percentage of actual data in the respective month of the previous year	ttage data in ective he previ-	Forecast values (billion USD a month)		Percentage of actual data in the respective month of the previous year	ttage data in ective he previ-
	ARIMA	$_{ m SM}$	ARIMA	$_{ m SM}$	ARIMA	$_{ m SM}$	ARIMA	SM	ARIMA	SM	ARIMA	SM	ARIMA	$_{ m SM}$	ARIMA	SM
Sep 16	24.1	23.2	91	88	14.9	15.6	88	93	20.9	19.5	93	98	14.0	13.5	93	06
Oct 16	22.5	25.8	83	96	17.1	16.4	101	97	21.2	21.9	93	96	15.3	14.3	102	95
Nov 16	21.7	26.3	85	104	15.9	16.7	26	101	20.0	21.6	93	101	14.4	12.8	86	87
Dec 16	23.7	28.6	85	66	16.8	16.5	96	95	20.3	22.4	83	91	14.7	15.5	94	66
Jan 17	19.0	18.0	111	105	9.3	10.5	94	107	15.8	15.5	107	105	0.6	8.6	102	97
Feb 17	21.6	19.8	108	66	12.9	14.1	100	109	18.7	17.4	108	100	12.9	11.0	111	95
				Н	Por referen	ce: actual	For reference: actual values in respective months of 2015/2016 (billion USD)	espective n	onths of 2	015/2016	billion US	()				
Sep 15		26.3	ç٠.			16.8	8.			21.3	ಣ			22.8	∞.	
Oct 15		27.0	0.			16.9	6.3			22.6	9.			24.0	0.	
Nov 15		25.4	.4			16.5	5.5			22.7	7			20.5	ಸರ	
Dec 15		28.7	7.			17.4	.4			21.5	ಬ			22.1	-	
Jan 16		17.1	.1			9.8	∞.			24.5	ಸರ			10.7	7	
Feb 16		20.1	.1			12.9	6:			14.8	∞.			13.9	6	

Note. Over the period from January 1999 to July 2016, the series of export, import, export to the countries outside the CIS and import from the countries outside the CIS were identified as stationary series in the first-order differences. In all the cases, seasonal components were included in the specification of the models.

#### FOREIGN TRADE INDICES

amounts to around 3.2% for September 2016 – February 2017 against the corresponding period of 2015–2016.

The average forecast decrease of the monthly real trade turnover for the period from September 2016 – February 2017 against the same period of 2015–2016 constitutes 3.4%.

At an annual rate, forecast growth of the nominal index of retail trade turnover in 2016 will come to 7.9%, and in real terms will decrease by 4.6%.

### **FOREIGN TRADE INDICES**

Model calculations of forecast values of the export and export to countries outside the CIS and the import and import from countries outside the CIS were made on the basis of the models of time series and structural models evaluated on the basis of the monthly data over the period from September 1998 to July 2016 on the basis of the data released by the Central Bank of Russia<sup>1</sup>. The results of calculations are shown in Table 3.

The average forecast decrease of exports, imports, exports outside the CIS and imports from the countries outside the CIS for September 2016 – February 2017 against the same period of 2015–2016 will amount to 5.1%, 2.2%, 4.7%, and 3.5%, respectively. The average forecast surplus volume of the trade balance with all countries for 2016 will constitute \$98.5bn, which corresponds to a decrease by 33.6% on the same period of 2015.

### DYNAMICS OF PRICES

### The Consumer Price Index and Producer Price Index

This section presents calculations of forecast values of the consumer price index and producer price index (as regards both the industry in general and some types of its activities under the National Industry Classification Standard (NICS)) made on the basis of the time-series models evaluated on the basis of the data released by Rosstat over the period from January 1999 to June 2016<sup>2</sup>. Table 4 presents the results of model calculations of forecast values over September 2016 and February 2017 in accordance with ARIMA models, structural models (SM) and models computed with the help of business surveys (BS).

The forecast average monthly growth of the consumer price index in September 2016 – February 2017 will come to 0.7%. The price growth of industrial goods for this period is forecast at an average monthly rate of 0.2%. Annual growth of the consumer price index on average across three models will come to 6.5%. The same indicator for the producer price index is forecast at 9.0%.

For the producer price indices computed by Rosstat from September 2016 through February 2017 the following average monthly growth rates are forecast: for mining and quarrying (-0.2%), manufacturing 0.5%, utilities (electricity, water, and gas) 0.8%, food products 0.6%, textile and sewing industry 0.7%, wood products 0.5%, pulp and paper industry 0.5%, coke and refined petroleum 0.8%, for chemical industry 0.2%, for basic metals and fabricated metal 0.7%, for machinery and equipment 1.2%, and for transport equipment and manufacturing 0.5%.

Annual growth of the producer price indices across types of economic activity will average 10.1%. By end-2016, maximum annual growth is forecast in the production of basic metals and fabricated metal (22.7%), and the minimum – in chemical production (-0.3%).

<sup>1</sup> The data on the foreign trade turnover is calculated by the CBR in accordance with the methods for making of the balance of payment in prices of the exporter-country (FOB) in billion USD.

<sup>2</sup> Structural models were evaluated in the period from October 1998.

CALCULATIONS OF FORECAST VALUES OF PRICE INDICES

		for transport equipment manufacturing		100.4	100.7	101.0	100.2	101.1	8.66		105.4	106.2	107.2	107.5	101.1	100.9		116.1	116.6	115.9	116.3	6.66	101.0
		for machinery and equip- ment		100.9	100.7	101.1	101.1	101.5	101.7		109.6	110.4	111.6	112.8	101.5	103.2		111.6	113.0	113.5	114.0	103.3	104.0
		-dsi bas sletans sad rot ricated metal		101.3	100.6	100.3	100.8	100.5	100.6		120.5	121.3	121.7	122.7	100.5	101.1		114.3	114.7	113.1	111.9	99.7	102.3
		for chemical industry		100.1	100.1	100.1	100.3	100.4	100.4		99.3	99.4	99.4	99.7	100.4	100.7		118.6	119.9	119.5	118.2	100.2	101.9
		for coke and refined pet-		101.5	100.9	100.5	100.8	100.8	100.4		113.9	115.0	115.5	116.5	100.8	101.2	15)	114.3	113.3	108.7	100.9	92.6	92.7
		rog pand paper Yatsubni		100.5	100.6	100.1	100.5	100.6	100.6		109.7	110.4	110.5	111.2	100.6	101.2	r 2014/2015)	115.7	116.9	118.2	119.0	101.2	106.1
	ndices:	for wood products		100.3	100.2	100.1	100.5	100.8	101.0		103.7	103.9	104.0	104.5	100.8	101.8	of 2015/2016 (% of December	109.4	109.2	109.3	109.3	100.5	100.2
	Producer price indices:	gniwəs bns əlitxət rof yrtsubni	month)	100.7	100.7	100.8	100.3	100.8	100.7	15/2016)	108.6	109.3	110.2	110.6	100.8	101.5	016 (% of	111.9	112.9	113.6	113.6	102.0	103.3
, ,EO L	Produc	stoubord boot rot	orecast values (% of the previous month)	100.3	100.4	100.4	100.8	100.7	101.2	(% of December 201	104.7	105.1	105.6	106.4	100.7	101.9	of 2015/2	111.2	112.1	113.2	113.9	100.8	101.3
- ()		for utilities (electricity, water, and gas)	(% of the	100.3	100.2	100.9	100.2	101.1	102.0	(% of Dec	102.1	102.3	103.2	103.4	101.1	103.0	e periods	105.8	107.9	109.2	109.4	99.3	100.5
5		garrutəsfunsm rot	st values	100.9	100.6	8.66	6.66	100.8	100.9	recast values	109.5	110.2	109.9	109.8	100.8	101.7	ues in the same periods	113.0	113.5	112.7	111.2	99.4	100.6
		garyraeup bas gaiaim rof	Foreca	8.76	101.2	100.6	0.86	100.6	100.8	Foreca	116.1	117.5	118.1	115.8	100.6	101.4	valı	112.7	120.0	118.0	109.8	92.6	83.8
() (-) (-)		(MH) shoog Isirtsubni rol		101.6	100.4	100.7	8.66	6.66	101.3		110.9	111.3	112.1	111.9	6.66	101.2	For reference: actual						
		(SB) shoog Isirtsubni rol		99.4	100.8	99.4	2.66	9.66	100.4		105.2	106.0	105.4	105.1	9.66	100.0	or referen	112.0	114.0	113.2	110.8	98.6	97.0
		sboog Isirtsubni rof (AMIAA)		100.3	100.8	99.5	100.2	100.1	100.2		109.4	110.3	109.8	110.1	100.1	100.3	Fc						
	X	The consumer price index (FM)		100.4	100.5	100.7	100.7	101.8	100.8		104.7	105.2	106.0	106.7	101.8	102.6							
	X	The consumer price index (MS)		100.1	100.3	100.4	100.4	100.6	100.5		104.0	104.3	104.7	105.1	100.6	101.1		110.6	111.4	112.3	113.2	101.0	101.6
	X	The consumer price index (AMIAA)		100.6	100.8	100.9	100.9	101.5	101.1		104.8	105.7	106.6	107.6	101.5	102.6							
				Sep 16	Oct 16	Nov $16$	$\mathrm{Dec}\ 16$	Jan 17	Feb 17		Sep 16	Oct 16	Nov 16	Dec 16	Jan 17	Feb 17		Sep 15	Oct 15	Nov $15$	$\mathrm{Dec}\ 15$	Jan 16	Feb 16

**Note.** Over the period from January 1999 to June 2016, the series of the chain producer price index for machinery are identified as a stationary process around the trend with two endogenous structural changes. The series of other chain price indices are stationary at levels.

#### Table 5

### The Cost of the Monthly per Capita Minimum Food Basket

This section presents calculations of forecast values of the cost of the monthly per capita minimum food basket over September 2016 and February 2017. The forecasts were made based on time series with use the Rosstat data over the period from January 2000 to July 2016. The results are shown in Table 5.

As can be seen from *Table 5*, cost growth of the minimum set of food products is forecast compared to the corresponding period of the previous year. Herewith, forecast cost of the minimum set of food products constitutes around RUR 3,764.8. The forecast cost growth of the minimum set of food products will average around 5.3% compared to the level of the corresponding period of the previous year. Annual growth of the cost of the minimum set of food products in 2016 will constitute 6%.

### **Indices of Freight Rates**

This section presents calculations of forecast values of freight rate indices on cargo carriage<sup>1</sup>, made on the basis of time-series models evaluated on the Rosstat data over the period from September 1998 to July 2016. Table 6 shows the results of model calculations of forecast values in September 2016 - February 2017. It should be noted that some of the indices under review (for instance, the pipeline rate index) are adjustable ones and for that reason their behavior is hard to describe by means of the time-series models. As a result, the future values may differ greatly from the real ones in case of the centralized increase of rates in the period of forecasting or in case of absence of such an increase in the forecasting period, but with it taking place shortly before the beginning of that period.

According to the forecast results for September 2016 – February 2017, the composite freight rate index will decrease on average 0.2% per month. As a result, its annual growth in 2016 will come to 10.5%.

### THE FORECAST OF THE COST OF THE MONTHLY PER CAPITA MINIMUM FOOD BASKET

Forecast values acc	cording to ARIMA-model (RUR)
Sep 16	3,600.7
Oct 16	3,625.7
Nov 16	3,697.5
Dec 16	3,804.0
Jan 17	3,894.4
Feb 17	3,966.1
	ual values in the same months
	/2016 (billion RUR)
Sep 15	3,516.7
Oct 15	3,516.5
Nov 15	3,547.2
Dec 15	3,589.9
Jan 16	3,627.1
Feb 16	3,649.8
Expected grow	th on the respective month
of the	previous year (%)
Sep 16	2.4
Oct 16	3.1
Nov 16	4.2
Dec 16	6.0
Jan 17	7.4
Feb 17	8.7

**Note.** The series of the cost of the monthly per capita minimum food basket over the period from January 2000 to July 2016 are stationary in the first-order differences.

Table 6

### CALCULATIONS OF FORECAST VALUES OF INDICES OF FREIGHT RATES

		I INLIGITINA	ILJ
Period	The composite freight rate index	The index of truckload freight rate	The index of pipeline rate
I		ecording to ARIMA	A-models
~		previous month)	1000
Sep 16	99.8	99.9	100.2
Oct 16	99.8	100.0	97.8
Nov 16	99.8	99.9	102.0
Dec 16	99.8	99.9	103.3
Jan 17	99.8	101.6	99.3
Feb 17	99.8	99.9	98.6
I	Forecast values a	ccording to ARIMA	A-models
	(% of December	er of the previous	year)
Sep 16	111.2	100.6	116.6
Oct 16	111.0	100.6	114.0
Nov 16	110.8	100.5	116.3
Dec 16	110.5	100.5	120.1
Jan 17	99.8	101.6	99.3
Feb 17	99.6	101.5	97.9
Fo	or reference: actua	al values in the sa	me period
	of 2015/2016 (%	of the previous n	nonth)
Sep 15	100.1	100.5	99.9
Oct 15	94.5	99.6	89.6
Nov 15	100.2	100.2	100.2
Dec 15	100.6	101.9	100.5
Jan 16	100.7	102.2	93.7
Feb 16	99.8	100.1	99.8

Note. Over the period from September 1998 to July 2016, the series of the freight rates index were identified as stationary ones; the other series were identified as stationary ones over the period from September 1998 to July 2016, too; fictitious variables for taking into account particularly dramatic fluctuations were used in respect of all the series.

<sup>1</sup> The paper presents a review of the composite freight rate index on freight transport and the truckload freight rate index, as well as the pipeline rate index. The composite freight rate index is computed on the basis of the freight rate indices by individual types of transport: rail, pipeline, shipping, domestic water-borne, and truckload freight and air service (for more detailed information, pls. refer, for instance, to: *Prices in Russia*. The Official Publication of Goskomstat of RF, 1998).

### 8'2016 MODEL CALCULATIONS OF SHORT-TERM FORECASTS...

The index of truckload freight rate will grow at a monthly average rate of 0.2% in the course of given six months. Its annual growth is forecast at 0.5% in 2016.

Pipeline transport index will be growing in the course of next six months at a monthly average rate of 0.2%. As a result, its annual growth will amount to 20.1% in 2016.

### **World Prices of Natural Resources**

This section presents calculations of such average monthly values of Brent crude prices (US\$ per barrel), the aluminium prices (US\$ per ton), the gold prices (\$ per ounce), the copper prices (US\$ per ton) and the nickel prices (US\$ per ton) over September 2016 and February 2017 as were received on the basis of nonlinear models of time series evaluated on the basis of the IMF data over the period from January 1980 to July 2016.

Table 7
CALCULATIONS OF FORECAST VALUES OF WORLD PRICES ON NATURAL RESOURCES

G.	Brent oil	Aluminum	Gold	Copper	Nickel
Month	(\$ per barrel)	(\$ per ton)	(\$ per ounce)	(\$ per ton)	(\$ per ton)
	(φ per barrer)	(1)	cast values	(φ per ton)	(\$ per ton)
9 10					
Sep 16	40.06	1,658	1,371	5,003	11,041
Oct 16	46.63	1,675	1,379	5,047	11,325
Nov 16	42.25	1,672	1,398	5,080	11,537
Dec 16	42.56	1,669	1,417	5,100	11,664
Jan 17	43.12	1,670	1,431	5,118	11,869
Feb 17	48.33	1,662	1,439	5,136	11,942
	Expected	growth on the respec	ctive month of the pro	evious year (%)	
Sep 16	-15.2	4.3	21.9	-4.1	11.1
Oct 16	-3.1	10.5	19.0	-3.2	9.8
Nov 16	-4.9	13.9	28.8	5.8	24.8
Dec 16	12.8	11.5	32.6	9.9	34.0
Jan 17	40.0	12.8	30.4	14.4	39.5
Feb 17	45.6	8.6	20.0	11.7	43.9
	For ref	ference: actual values	s in the same period o	of 2015/2016	
Sep 15	47.23	1,590	1,125	5,217	9,938
Oct 15	48.12	1,516	1,159	5,216	10,317
Nov 15	44.42	1,468	1,086	4,800	9,244
Dec 15	37.72	1,497	1,068	4,639	8,708
Jan 16	30.8	1,481	1,097	4,472	8,507
Feb 16	33.2	1,531	1,200	4,599	8,299

**Note.** Over the period from January 1980 to July 2016, the series of prices of crude oil, nickel, gold, copper and aluminum are series of DS type.

The average forecast of crude oil price amounts to around \$43.8 per barrel, which is above its corresponding year-earlier indices on average by 12.5%. Aluminum prices are forecast at around \$1,668.0 per ton and their average forecast growth constitutes around 10.0% compared to the same level last year. Forecast for gold prices constitute around \$1,406.0 per ounce. Forecast average copper prices constitute around \$5,081 per ton and of nickel prices — about \$11,563 per ton. Average forecast price growth on gold constitutes around 25%, average reduction of copper prices — about 6%, average reduction of nickel prices — 27% compared to the corresponding level of last year.

By end-2016, forecast growth of prices on crude oil, aluminum, gold, copper and nickel against end-2015 will come to 12.8%, 11.5%, 32.7%, 9.9%, and 33.9% respectively.

### **MONETARY INDICES**

The future values of the monetary base (in the narrow definition: cash funds and the Fund of Mandatory Reserves (FMR) and M<sub>o</sub> monetary aggregate over the period from September 2016 to

February 2017 were received on the basis of models of time-series of respective indices calculated by the CBR¹ over the period from October 1998 to August (July - for M2 time series) 2016. Table 8 presents the results of calculations of forecast values and actual values of those indices in the same period of previous year. It is to be noted that due to the fact that the monetary base is an instrument of the CBR policy, forecasts of the monetary base on the basis of time-series models are to a certain extent notional as the future value of that index is determined to a great extent by decisions of the CBR, rather than the inherent specifics of the series.

In September 2016 – February 2017, the monetary base will be growing at an average monthly rate of 0.8%, and the monetary indicator  $M_2$  – at an average monthly rate of 0.9%. In 2016, annual growth of the indicator M2 is forecast at the level of 12.5%, and the monetary base -10.8%.

### INTERNATIONAL RESERVES

This section presents the outputs of the statistical estimation of such future values of the international reserves of the Russian Federation<sup>2</sup> as were received on the basis of evaluation of the model of time series of the gold and foreign exchange reserves on the basis of the data released by the CBR over the period from October 1998 to August 2016. That index is forecast without taking into account a decrease in the amount of reserves due to foreign debt payment and for that reason the values of the volumes of the international reserves in the months where foreign debt payments are made may happen to be overestimated (or, otherwise, underestimated) as compared to the actual ones.

Subsequent to the forecast results in September 2016 – February 2017, the international reserves will be growing by an average monthly rate of 0.5%. As a result, annual growth of the international reserves in 2016 will come to 10.3%.

THE FORECAST OF Ma AND THE MONETARY BÂSE

	The M	onetary base		${ m M}_2$
Period	Billion RUR	Growth on the previous month, %	Billion RUR	Growth on the previous month, %
Sep 16	8,722	2.0	36,821	0.5
Oct 16	8,666	-0.6	36,987	0.5
Nov 16	8,852	2.1	37,153	0.4
Dec 16	8,798	-0.6	37,477	0.9
Jan 17	9,321	6.0	38,548	2.9
Feb 17	8,931	-4.2	38,714	0.4
		ctual value in		

of 2015/2016 (growth on the previous month, %)

Sep 15	0.6	1.1
Oct 15	-1.6	-0.2
Nov 15	0.2	-0.3
Dec 15	-0.1	1.4
Jan 16	10.1	7.5
Feb 16	-6.3	-2.7

Note. Over the period from October 1998 to August (July) 2016, all the time series of monetary indices were attributed to the class of series which are stationary in the first-order differences and have an explicit seasonal component.

Table 9 THE FORECAST OF THE INTERNATIONAL RESERVES OF THE RUSSIAN FEDERATION

	Forecast valu	ues according to ARIMA-model			
Period	Billion USD	Growth on the previous			
	Dillion CDD	month, %			
Sep 16	396.2	0.6			
Oct 16	399.4	0.8			
Nov 16	400.8	0.3			
Dec 16	402.3	0.4			
Jan 17	404.2	0.5			
Feb 17	406.0	0.5			
For 1	reference: actu	al values in the same period			
	of	f 2015/2016			
Sep 15	366.3	2.4			
Oct 15	371.3	1.3			
Nov 15	369.6	-0.4			
Dec 15	364.7	-1.3			
Jan 16	368.4	1.0			
Feb 16	371.6	0.9			

Note. Over the period from October 1998 to August 2016, the series of the gold and foreign exchange reserves of the Russian Federation were identified as stationary series in difference.

The data on the specific month is given in accordance with the methods of the CBR as of the beginning of the following

The data on the volume of the gold and foreign exchange reserves is presented as of the first day of the following month.

### **FOREIGN EXCHANGE RATES**

The model calculations of prospective values of the foreign exchange rates (RUR per USD and USD per euro) were made on the basis of assessment of the time series models (ARIMA) and structural models (SM) of the relevant indicators released by the Central Bank of Russia as of the last date of each month over the periods from October 1998 to August 2016 and from January 1999 to August 2016<sup>1</sup>, respectively.

USD/RUR average exchange rate in the intervening period is forecast on average along two models in the amount of RUR 65.77 for USD. Forecast by end-2016 average (along two models) indicator will total Rb 66.19 for USD.

Euro/USD average exchange rate is forecast at USD 1.11 per 1 euro. By end-2016, the indicator is forecast at USD 1.12 per 1 euro along two models.

### THE LIVING STANDARD INDICES

This section (Table 12) presents calculations of forecast values of indices of real wages, real disposable income and real income<sup>2</sup> as were received on the basis of the model of time series of respective indices computed by Rosstat and taken over the period from January 1999 to July 2016. The above indices depend to a certain extent on the centralized decisions on raising of wages and salaries to public sector workers, as well as those on raising of pensions, scholarships and allowances; such a situation introduces some changes in the dynamics of the indices under review. As a result, the future values of the indices of real wages and real disposable income calculated on the basis of the series which last observations are either considerably higher or lower than the previous ones due to such a raising may differ greatly from those which are implemented in reality.

According to the results presented in *Table 11*, the forecast average monthly fall of the real disposable cash income will constitute 2.7% per month

FORECASTS OF THE USD/RUR AND EUR/USD EXCHANGE RATES

	exchan	SD/RUR ge rate er USD)	The EU exchan (USD pe	ge rate		
	ARIMA	SM	ARIMA	SM		
Sep 16	1.11	1.10	65.45	65.99		
Oct 16	1.11	1.12	65.03	64.74		
Nov 16	1.11	1.11	65.47	65.89		
Dec 16	1.11	1.11	65.75	66.10		
Jan 17	1.11	1.12	66.06	66.39		
Feb 17	1.11	1.13	66.37	66.00		
For re	eference: ac		in the simil	ar period		
		of 2015/20				
Sep 15	66.	.24	1.	12		
Oct 15	64.	.37	1.	10		
Nov 15	66.	.24	1.0	05		
Dec 15	72.	.88	1.0	09		
Jan 16	75.	.17	1.0	09		

**Note.** Over the respective periods, the series under review were identified as integrated series of the first order with a seasonal component.

1.09

75.09

Feb 16

Table 11
THE FORECAST OF THE LIVING STANDARD
INDICES

	Real dispos- able cash income	Real cash income	Real accrued wages
For	ecast values accor	ding to ARII	MA-models
(9	% of the respective	month of 20	015/2016)
Sep 16	96.4	96.2	103.2
Oct 16	95.9	95.7	103.2
Nov 16	97.7	97.6	104.8
Dec 16	96.3	95.9	104.2
Jan 17	99.2	98.5	103.7
Feb 17	98.1	98.1	100.7
For	reference: actual	values in the	erespective
	period of	2015/2016	
	(% of the same pe	eriod of 2014	/2015)

	period of	2015/2016	
	(% of the same pe	eriod of 2014	(2015)
Sep 15	93.9	93.8	89.6
Oct 15	93.2	93.2	89.5
Nov 15	93.7	93.5	89.6
Dec 15	99.1	98.5	91.6
Jan 16	94.2	95.0	96.4
Feb 16	95.7	95.3	100.6
Feb 16	95.7	95.3	100.6

Note. For calculating purposes, the series of the real disposable cash income, real cash income and real accrued wages in the base form were used (January 1999 was adopted as a base period). Over the period from January 1999 to July 2016 those series were attributed to the class of processes which are stationary in differences and have an explicit seasonal component.

<sup>1</sup> The authors use the IMF data over the period from January 1999 to July 2016. The data over the period from July and August 2016 was obtained from the foreign exchange rate statistics website: www.oanda.com

<sup>2</sup> Real cash income is a relative index which is calculated by means of division of the index of the nominal size (which was actually formed in the period under review) of households' cash income by the CPI. Real disposable cash income is cash income minus mandatory payments and contributions. (See: Rossiisky Statistichesky Ezhegodnik, Moscow, Rosstat, 2004, p. 212).

### **FMPI OYMENT AND UNEMPI OYMENT**

compared to the same period of last year; the real cash income – 3.0%. The average monthly growth of the real accrued wages is forecast in the amount of 3.3% in comparison with the same period of the previous year.

By end-2016, forecast decrease of the real disposable cash income will amount to 3.4%; the real cash income – by 4.5%, and growth of the level of the real wages – 1.4%.

### **EMPLOYMENT AND UNEMPLOYMENT**

For the purpose of calculation of the future values of the employment (of the number the gainfully employed population) and the unemployment (the total number of the unemployed), models of the time series evaluated over the period from October 1998 to June 2016 on the basis of the monthly data released by Rosstat<sup>1</sup> were used. The unemployment was calculated on the basis of the models with results of the findings from business surveys<sup>2</sup>, too.

It is to be noted that feasible logical inconsistencies<sup>3</sup> in forecasts of employment and unemployment which totals should be equal to the index of economically active population may arise due to the fact that each series is forecast individually and not as a difference between the forecast values of the economically active population and another index.

Table 12
CALCULATION OF FORECAST VALUES OF THE INDICES THE EMPLOYMENT AND THE UNEMPLOYMENT

CAL	COLATIC	IN OF FORECE	VALO	L3 OI IIIL IIAD	ICLS IIIL LIVII L	OTATEIN	AND THE UNER	VII LO I /VILINI
	Employ	ment (ARIMA)		Unemployment (A	ARIMA)		Unemployment	(BS)
	Million people	Growth on the respective month of previ- ous year (%)	Million people	Growth on the respective month of previ- ous year (%)	% of the index of the number of the gainfully employed popu- lation	Million people	Growth on the respective month of previ- ous year (%)	% of the index of the number of the gainfully employed popu- lation
Sep 16	73.2	0.4	4.2	5.3	5.8	4.2	3.8	5.7
Oct 16	72.8	0.5	4.3	0.9	6.0	4.4	1.3	6.0
Nov 16	72.6	0.6	4.4	0.4	6.1	4.4	1.0	6.1
Dec 16	72.4	0.2	4.5	2.6	6.2	4.5	1.2	6.2
Jan 17	72.0	1.0	4.7	6.5	6.5	4.5	1.4	6.3
Feb 17	72.0	0.7	4.8	8.1	6.6	4.5	2.1	6.3
		For referen	ce: actual	values in the sam	ne periods of 2015	/2016 (mil	lion people)	
Sep 15		72.9				4		
Oct 15		72.5			4	.3		
Nov 15		72.2			4	.4		
Dec 15		72.3			4	.4		
Jan 16		71.3			4	.4		
Feb 16		71.5			4	.4		

**Note.** Over the period from October 1998 to June 2016, the series of employment is a stochastic process which is stationary around the trend. The series of unemployment is a stochastic process with the first order integration. Both indices include seasonal component.

According to ARIMA-model forecast (*Table 12*), in September 2016 – February 2017, the decrease of the number of employed in the economy on average will constitute 0.5% per month against the corresponding period of the previous year. Forecast by end-2016 indicator of the number of employed in the economy constitutes 72.4 mn persons.

The average increase of the total number of jobless is forecast at 2.9% per month against the corresponding period last year. Average number of jobless by end-2016 is forecast at 4.5 mn persons.

<sup>1</sup> The index is computed in accordance with the methods of the International Labor Organization (ILO) and is given as of the month-end.

<sup>2</sup> The model is evaluated over the period from January 1999 to June 2016.

<sup>3</sup> For example, deemed as such a difference may be a simultaneous decrease both in the employment and the unemployment. However, it is to be noted that in principle such a situation is possible provided that there is a simultaneous decrease in the number of the economically active population.

### **ANNEX**

### Diagrams of the Time Series of the Economic Indices of the Russian Federation

Fig. 1a. The Rosstat industrial production index (ARIMA-model) (% of December 2001)

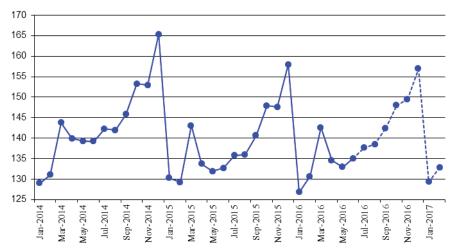


Fig. 1b. The NRU HSE industrial production index (ARIMA-model) (% of January 2005)

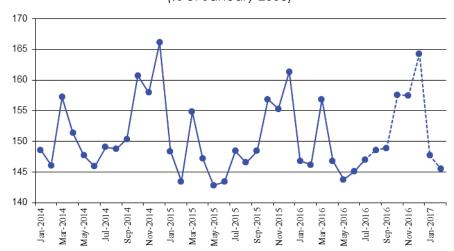


Fig. 2a. The Rosstat industrial production index for mining (% of December 2001)

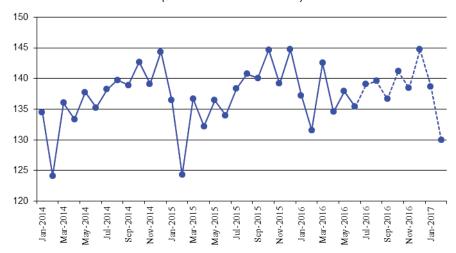


Fig. 2b. The NRU HSE industrial production index for mining (% of January 2005)

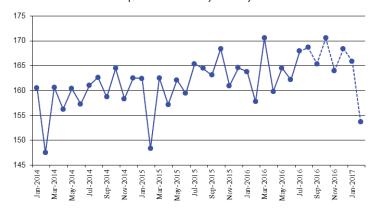


Fig. 3a. The Rosstat industrial production index for manufacturing (% of December 2001)



Fig. 3b. The NRU HSE industrial production index for manufacturing (% of January 2005)



Fig. 4a. The Rosstat industrial production index for utilities (electricity, water, and gas) (as a percentage of that in December 2001)

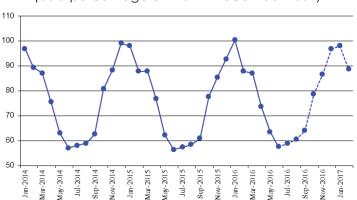


Fig. 4b. The NRU HSE industrial production index for utilities (electricity, water, and gas) (as a percentage of that in January 2005)

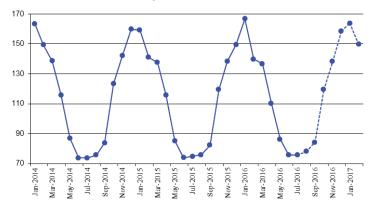


Fig. 5a. The Rosstat industrial production index for food products (as a percentage of that in December 2001)

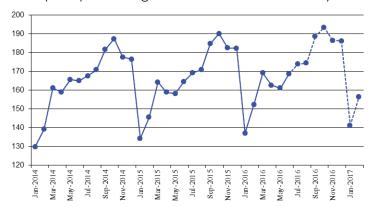


Fig. 5b. The NRU HSE industrial production index for food products (as a percentage of that in January 2005)

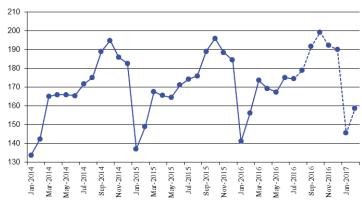


Fig. 6a. The Rosstat industrial production index for coke and petroleum (as a percentage of that in December 2001)

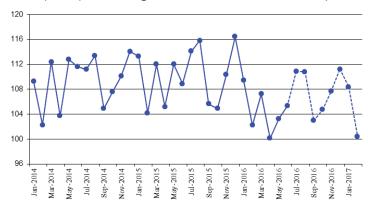


Fig. 6b. The NRU HSE industrial production index for petroleum and coke (as a percentage of that in January 2005)

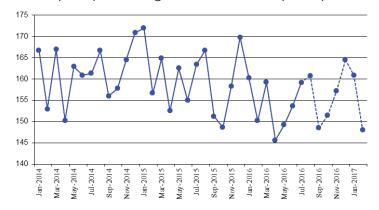


Fig.7a. The Rosstat industrial production index for primary metals and fabricated metal products (as a percentage of that in December 2001)

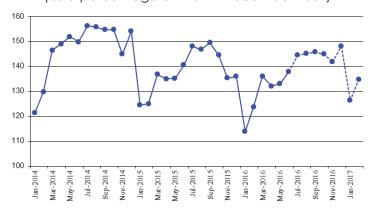


Fig. 7b. The NRU HSE industrial production index for primary metals and fabricated metal products (as a percentage of that in January 2005)

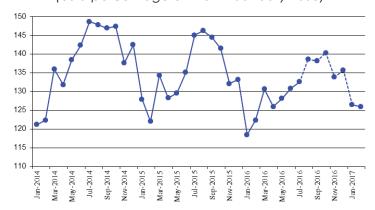


Fig. 8a. The Rosstat industrial production index for machinery (as a percentage of that in December 2001)

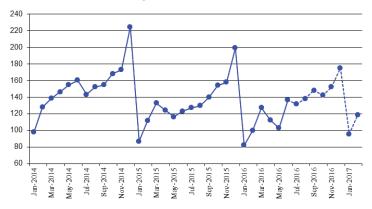


Fig. 8b. The NRU HSE industrial production index for machinery (as a percentage of that in January 2005)

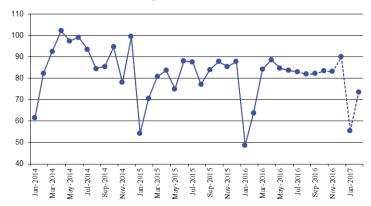


Fig. 9. The volume of retail sales (billion RUR)

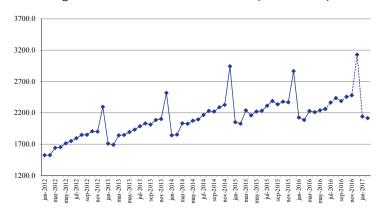


Fig. 9a. The real volume of retail sales (as a percentage of that in the same period of the previous year)

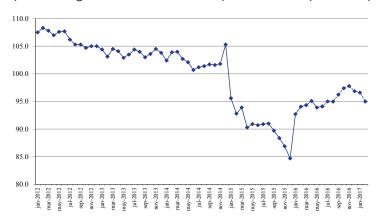


Fig. 10. Export to all countries (billion USD)

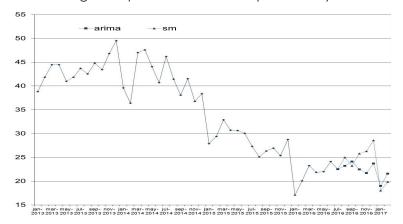


Fig. 11. Export to countries outside the CIS (billion USD)

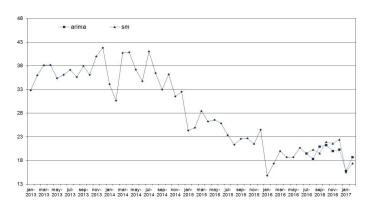


Fig. 12. Import from all countries (billion USD)

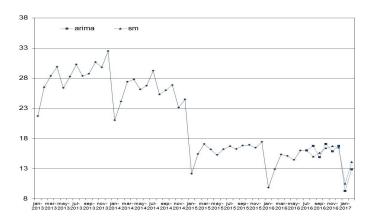


Fig. 13. Import from countries outside the CIS (billion USD)

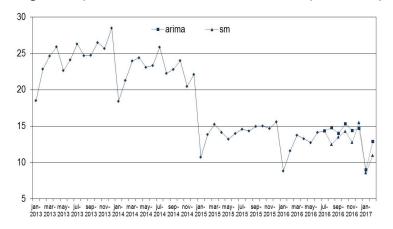


Fig. 14. The consumer price index (as a percentage of that in December of the previous year)

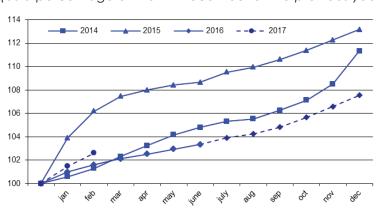


Fig. 14a. The consumer price index (as a percentage of that in December of the previous year) (SM)

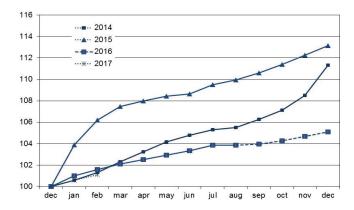


Fig.15. The producer price index for industrial goods (as a percentage of that in December of the previous year)

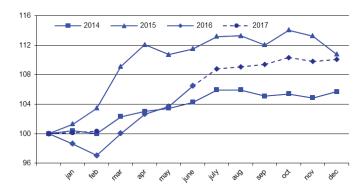


Fig. 16. The price index for mining (as a percentage of that in December of the previous year)

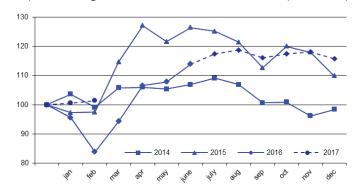


Fig. 17. The price index for manufacturing (as a percentage of that in December of the previous year)

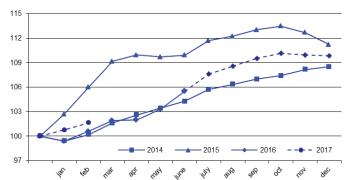


Fig. 18. The price index for utilities (electricity, water, and gas) (as a percentage of that in December of the previous year)

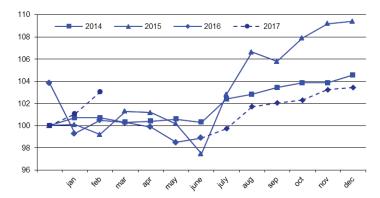


Fig. 19. The price index for food products (as a percentage of that in December of the previous year)

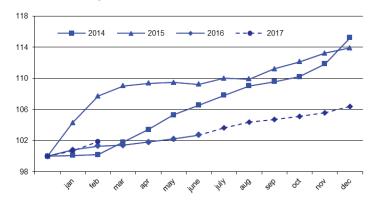


Fig. 20. The price index for the textile and sewing industry (as a percentage of that in December of the previous year)

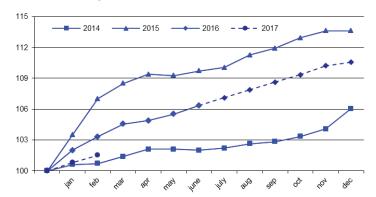


Fig. 21. The price index for wood products (as a percentage of that in December of the previous year)

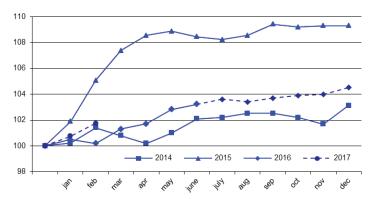


Fig. 22. The price index for the pulp and paper industry (as a percentage of that in December of the previous year)

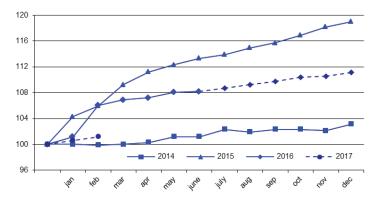


Fig. 23. The price index for coke and petroleum (as a percentage of that in December of the previous year)

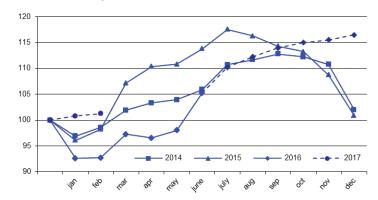


Fig. 24. The price index for the chemical industry (as a percentage of that in December of the previous year)

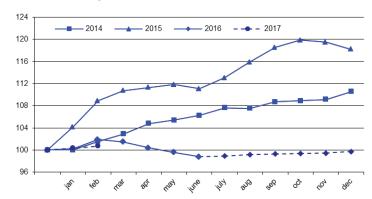


Fig.25. The price index for primary metals and fabricated metal products (as a percentage of that in December of the previous year)

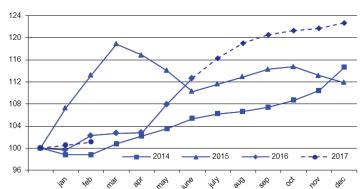


Fig.26. The price index for machinery (as a percentage of that in December of the previous year)

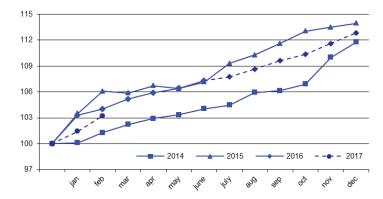


Fig.27. The price index for transport equipment manufacturing (as a percentage of that in December of the previous year)

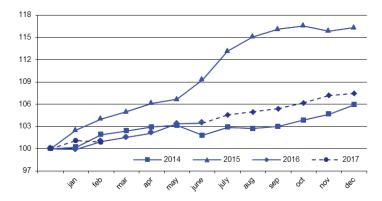


Fig. 28. The cost of the monthly per capita minimum food basket (RUR)

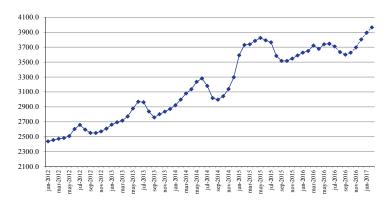
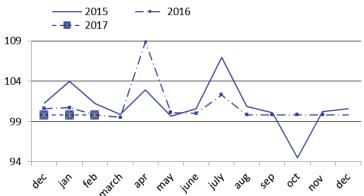


Fig. 29. The composite index of transport tariffs (for each year. as a percentage of that in the previous month)



### 8'2016 MODEL CALCULATIONS OF SHORT-TERM FORECASTS...

Fig. 30. The index of motor freight tariffs (for each year, as a percentage of that in the previous month)

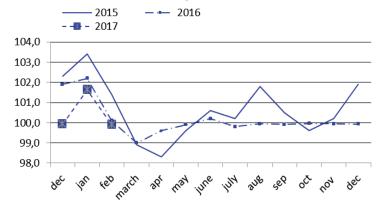


Fig. 31. The index of pipeline tariffs (for each year, as a percentage of that in the previous month)

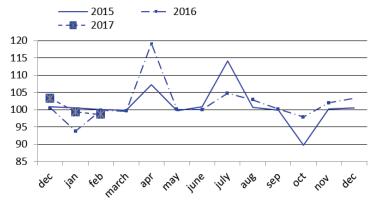


Fig. 32. The Brent oil price (\$ per barrel)

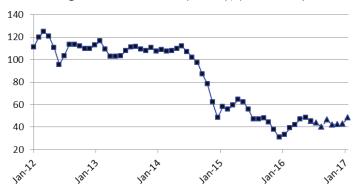
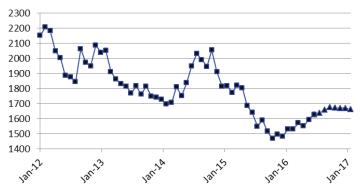
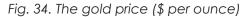


Fig. 33. The aluminum price (\$ per ton)





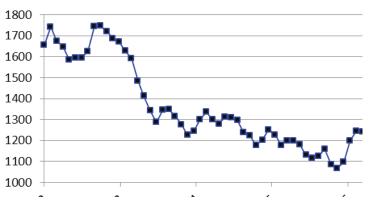


Fig. 35. The nickel price (\$ per ton)

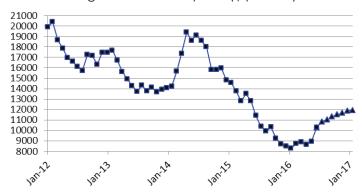


Fig. 36. The copper price (\$ per ton)

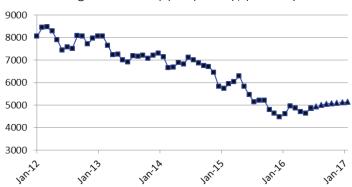
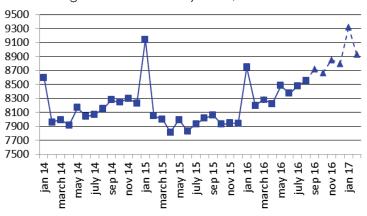


Fig. 37. The monetary base, billion RUR



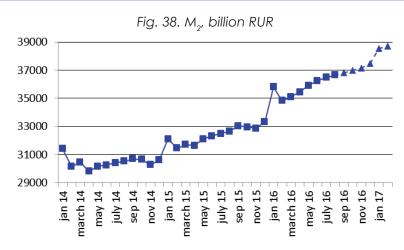


Fig. 39. The international reserves of the Russian Federation, million USD

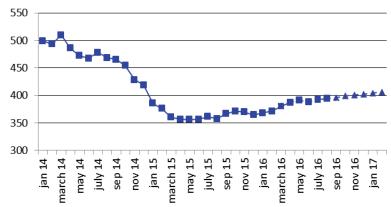


Fig. 40. The RUR/USD exchange rate

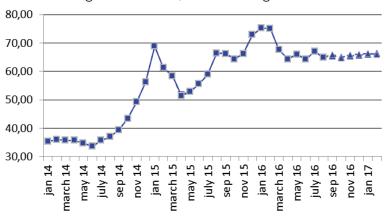


Fig. 41. The USD/EUR exchange rate

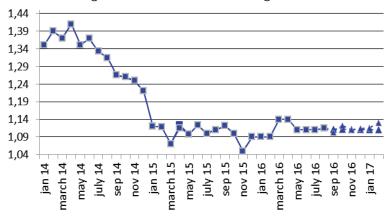


Fig. 42. Real disposable cash income (as a percentage of that in the same period of the previous year)

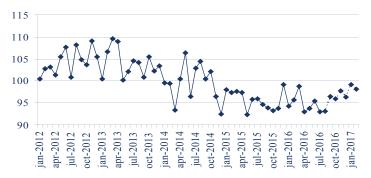


Fig. 43. Real cash income (as a percentage of that in the same period of the previous year)

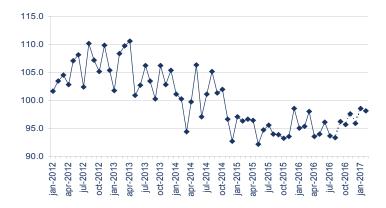


Fig. 44. Real accrued wages (as a percentage of those in the same period of the previous year)

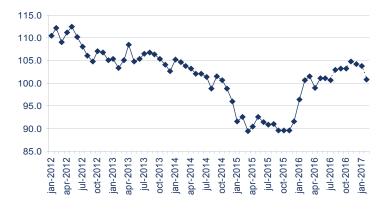
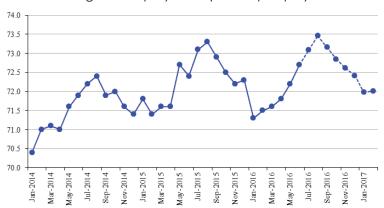


Fig. 45. Employment (million people)



### 8'2016 MODEL CALCULATIONS OF SHORT-TERM FORECASTS...

5.00 4.75 4.50 4.25 4.00

May-2015 Jul-2015 Sep-2015 Nov-2015 Mar-2016 May-2016

Jan-2016

Sep-2016

Jul-2016

Nov-2016

Jan-2017

3.50

Mar-2014

May-2014

Sep-2014

Nov-2014

Jul-2014

Jan-2015 Mar-2015

Fig. 46. Unemployment (million people)

Model calculations of short-term forecasts of social and economic indices of the Russian Federation: August 2016

MOUCH CAICULATIONS OF SHOT PICTIN TOLOGASES OF SOCIAL AND		cconomic marcs		of the forssian reactauting		Mugust 2010	_		
Index	June $9016$	July 2016	August $9016$	September	October $2016$	November 2016	December	January	February
Rosstat IIIP (growth rate, %)*	1.7	1.4	2.1	1.2	0.2	1.6	0.1	2.1	1.3
HSE IIP (growth rate %)*	1.1	6.0-	1.7	0.5	0.3	1.5	1.2	1.3	0.1
Rosstat IIP for mining (growth rate, %)*	1.6	0.5	8.0-	-2.4	-2.4	-0.5	0.0	1.1	-1.2
HSE IIP for mining (growth rate, %)*	1.7	1.6	2.5	1.4	1.3	1.9	2.3	1.3	-2.6
Rosstat IIIP for manufacturing (growth rate, %)*	1.6	1.5	4.0	-0.5	-1.1	6.0-	-2.4	-1.1	-3.1
HSE IIP for manufacturing (growth rate, %)*	9.0	-2.8	9.0	-1.5	-1.2	0.2	-2.2	0.1	-2.5
Rosstat IIP for utilities (electricity, water, and gas) (growth rate, $\%$ )*	2.0	2.5	3.6	5.2	1.2	1.5	4.4	-2.3	1.0
HSE for utilities (electricity, water, and gas) (growth rate, %)*	2.4	1.2	2.8	2.4	0.0	-0.1	6.0	-1.7	7.1
Rosstat IIP for food products (growth rate, %)*	2.3	2.6	2.1	2.1	1.8	2.2	2.2	3.1	2.7
HSE IIP for food products (growth rate, %)*	2.4	0.1	1.7	1.4	1.7	2.0	3.1	3.2	1.7
Rosstat IIP for coke and petroleum (growth rate, %)*	-1.8	-2.9	-4.3	-2.5	-0.2	-2.5	-4.5	-1.0	-1.7
HSE for coke and petroleum (growth rate, %)*	-0.9	-2.6	-3.6	-1.8	1.9	9.0-	-3.1	0.4	-1.5
Rosstat for primary metals and fabricated metal products (growth rate, %)*	-1.0	-2.4	-1.1	-2.5	0.4	4.8	8.9	10.9	8.8
HSE IIP for primary metals and fabricated metal products (growth rate, %)*	-3.1	-8.6	-5.2	-4.4	-0.8	1.3	1.9	6.8	3.0
Rosstat IIP for machinery (growth rate, %)*	11.0	3.4	6.5	6.0	-7.8	-3.4	-12.2	16.6	18.8
HSE IIP for machinery (growth rate %)*	-4.9	-5.3	6.3	-2.2	-4.9	-2.7	2.4	14.0	15.1
Retail sales, trillion Rb	2.26	2.36	2.44	2.39	2.46	2.48	3.13	2.14	2.12
Real retail sales (growth rate, %)*	-5.90	-5.00	-5.02	-3.77	-2.59	-2.22	-3.14	-3.39	-5.01
Investments in capital assets, trillion Rb	22.5	24.1	23.7	24.2	24.0	26.2	18.5	20.7	22.5
Real investments in capital assets (growth rate, %)*	19.5	19.3	20.2	21.6	20.8	21.4	15.7	18.1	19.5
Export to all countries (billion \$)	16.0	15.9	15.3	16.8	16.3	16.7	9.6	13.5	16.0
Export to countries outside the CIS (billion \$)	14.3	13.7	13.8	14.8	13.6	15.1	8.8	12.0	14.3
Import from all countries (billion \$)	0.4	0.5	0.2	0.4	9.0	9.0	0.7	1.3	8.0
Import from countries outside the CIS (billion \$)	2.7	2.1	1.4	0.4	0.7	-0.1	-0.1	-0.1	9.0
	5.6	3.1	1.1	-2.2	1.2	9.0	-2.0	9.0	8.0
PPI for industrial goods (growth rate, %)**	2.1	2.0	0.0	0.9	9.0	-0.2	-0.1	8.0	6.0
PPI for mining (growth rate, %)**	0.4	6.0	2.0	0.3	0.2	6.0	0.2	1.1	2.0
PPI for manufacturing (growth rate, %)**	0.5	6.0	0.7	0.3	0.4	0.4	0.8	0.7	1.2
PPI for utilities (electricity, water, and gas) (growth rate, %)**	8.0	0.7	0.7	0.7	0.7	8.0	0.3	8.0	0.7
PPI for food products (growth rate, %)**	0.4	0.4	-0.2	0.3	0.2	0.1	0.5	8.0	1.0
PPI for the textile and sewing industry (growth rate, %)**	0.1	0.5	0.5	0.5	9.0	0.1	0.5	9.0	9.0
PPI for wood products (growth rate, $\%$ )**	7.4	4.7	1.9	1.5	0.9	0.5	8.0	0.8	0.4
PPI for the pulp and paper industry (growth rate, %)**	-0.8	0.2	0.2	0.1	0.1	0.1	0.3	0.4	0.4
PPI for coke and petroleum (growth rate, %)**	4.4	3.2	2.3	1.3	9.0	0.3	8.0	0.5	9.0

Index	June 2016	July 2016	August 2016	September 2016	October 2016	November 2016	December 2016	January 2017	February 2017
PPI for the chemical industry (growth rate, %)**	8.0	0.4	8.0	6.0	0.7	1.1	1.1	1.5	1.7
PPI for primary metals and fabricated metal products (growth rate, %)**	0.1	1.1	0.4	0.4	0.7	1.0	0.2	1.1	-0.2
PPI for machinery (growth rate, %)**	3.75	3.71	3.63	3.60	3.63	3.70	3.80	3.89	3.97
PPI for transport equipment manufacturing (growth rate, $\%$ )***	0.2	-0.2	-0.1	-0.1	0.0	-0.1	-0.1	1.6	-0.1
The cost of the monthly per capita minimum food basket (thousand Rb)	0.0	4.7	2.9	0.2	-2.2	2.0	3.3	2.0-	-1.4
The composite index of transportation tariffs (growth rate, $\%$ )***	0.0	2.3	-0.2	-0.2	-0.2	-0.2	-0.2	-0.2	-0.2
The index of pipeline tariffs (growth rate, %)**	48.5	45.1	43.8	40.1	46.6	42.2	42.6	43.1	48.3
The index of motor freight tariffs (growth rate, %)**	1.59	1.63	1.64	1.66	1.68	1.67	1.67	1.67	1.66
The Brent oil price (\$ a barrel)	1.28	1.34	1.36	1.37	1.38	1.40	1.42	1.43	1.44
The aluminum price (thousand \$ a ton)	4.64	4.86	4.94	5.00	5.05	5.08	5.10	5.12	5.14
The gold price (thousand \$ per ounce)	8.9	10.3	10.8	11.0	11.3	11.5	11.7	11.9	11.9
The nickel price (thousand \$ a ton)	8.37	8.47	8.56	8.72	8.67	8.85	8.80	9.32	8.93
The copper price (thousand \$ a ton)	36.3	36.5	36.7	36.8	37.0	37.2	37.5	38.5	38.7
The monetary base (trillion Rb)	0.39	0.39	0.39	0.40	0.40	0.40	0.40	0.40	0.41
M2 (trillion Rb)	64.26	67.05	64.91	65.72	64.89	65.68	65.93	66.23	66.19
Gold and foreign exchange reserves (billion \$)	1.11	1.11	1.11	1.11	1.12	1.11	1.11	1.12	1.12
The RUR/USD exchange rate (rubles per one USD)	-4.6	-7.0	6.9-	-3.6	-4.1	-2.3	-3.7	-0.8	-1.9
The USD/EUR exchange rate (USD per one Euro)	-3.9	-6.4	-6.7	-3.8	-4.3	-2.4	-4.1	-1.5	-1.9
Real disposable cash income (growth rate, %)*	1.1	9.0	2.9	3.2	3.2	4.8	4.2	3.7	0.7
Real cash income (growth rate, %)*	72.7	73.1	73.5	73.2	72.8	72.6	72.4	72.0	72.0
Real accrued wages (growth rate, %)*	4.2	4.2	4.2	4.2	4.3	4.4	4.5	4.7	4.8
Employment (million people)	72.3	72.0	72.1	72.3	72.4	72.9	72.7	73.0	73.2
Unemployment (million people)	4.4	4.6	3.1	4.5	4.4	4.2	4.2	4.3	9.1

Note. Actual values are printed in the bold type \* % of the respective month of the previous year \*\* % of the previous month.