10,2016

MODEL CALCULATIONS OF SHORT-TERM FORECASTS OF RUSSIAN ECONOMIC TIME SERIES

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INTRODUCTION TO ALL THE ISSUES

This paper presents calculations of various economic indicators for the Russian Federation in the period from November 2016 to April 2017, which were performed using time series models developed as a result of research conducted by the Gaidar Institute over the past few years¹. A method of forecasting falls within the group of formal or statistical methods. In other words, the calculated values neither express the opinion nor expert evaluation of the researcher, rather they are calculations of future values for a specific economic indicator, which were performed using formal ARIMA models (p, d, q) given a prevailing trend and its, in some cases, significant changes. The presented forecasts are of inertial nature, because respective models rely upon the dynamics of the data registered prior to the moment of forecasting and depend too heavily on the trends, which are typical of the time series in the period immediately preceding the time horizon to be forecast. The foregoing calculations of future values of economic indicators for the Russian Federation can be used in making decisions on economic policy, provided that the general trends, which were seen prior to forecasting for each specific indicator, remain the same, i.e. prevailing long-term trends will see no serious shocks or changes in the future.

Despite that there is a great deal of data available on the period preceding the crisis of 1998, models of forecasting were analyzed and constructed using only the time horizon which followed August 1998. This can be explained by the findings of previous studies² which concluded, among other key inferences, that the quality of forecasts was deteriorated in most of the cases when the data on the pre-crisis period was used. Additionally, it currently seems incorrect to use even shorter series (following the crisis of 2008), because statistical characteristics of models based on such a short time horizon are very poor.

Models for the economic indicators in question were evaluated using standard methods of time series analysis. Initially, the correlograms of the studied series and their first differences were analyzed in order to determine the maximum number of delayed values to be included into the specifications of a model. Then, the results of analyzed correlograms served as the basis for testing all the series for weak stationarity (or stationarity around the trend) using the Dickey–Fuller test. In some cases, the series were tested for stationarity around the segmented trend using Perron and Zivot–Andrews tests for endogenous structural changes³.

The series were broken down into weak stationary, stationary near the trend, stationary near the trend with structural change or difference stationary, and then models, which corresponded to each type (regarding the levels and including, if necessary, the trend or segmented trend or differences), were evaluated. The Akaike and Schwartz information criteria, the properties of models' residuals (lack of autocorrelation, homoscedasticity and normality) and the quality of the in-sample-forecasts based on these models were used to choose the best model. Forecast values were calculated for the best of the models constructed for each economic indicator.

Additionally, the Bulletin presents future monthly values of the CPI, which were calculated using models developed at the Gaidar Institute, and volumes of imports/exports from/to all countries, which were calculated using structural models (SM). The forecast values based on the structural models may, in some cases, produce better results than ARIMA-models do, because structural models are constructed by adding information of the dynamics of exogenous variables. Besides, the

See, for example, R.M. Entov, S.M. Drobyshevsky, V.P. Nosko, A.D. Yudin. The Econometric Analysis of the Time Series of the Main Macroeconomic Indices. Moscow, IET, 2001; R.M. Entov, V.P. Nosko, A.D. Yudin, P.A. Kadochnikov, S.S. Ponomarenko. Problems of Forecasting of Some Macroeconomic Indices. Moscow, IET, 2002; V. Nosko, A. Buzaev, P. Kadochnikov, S. Ponomarenko. Analysis of the Forecasting Parameters of Structural Models and Models with the Outputs of the Polls of Industries. Moscow, IET, 2003; M.Yu. Turuntseva and T.R. Kiblitskaya, Qualitative Properties of Different Approaches to Forecasting of Social and Economic Indices of the Russian Federation. Moscow, IET, 2010.

³ See.: Perron, P. Further Evidence on Breaking Trend Functions in Macroeconomic Variables, *Journal of Econometrics*, 1997, 80, pp. 355–385; Zivot, E. and D.W.K. Andrews. Further Evidence on the Great Crash, the Oil-Price Shock, and Unit-Root Hypothesis. *Journal of Business and Economic Statistics*, 1992, 10, pp. 251–270.

INTRODUCTION TO ALL THE ISSUES

use of structural forecasts in making aggregated forecasts (i.e. forecasts obtained as average value from several models) may help make forecast values more accurate.

The dynamics of the Consumer Price Index was modeled using theoretical assumptions arising from the monetary theory. The following was used as explanatory variables: money supply, output volume, the dynamics of the ruble-dollar exchange rate, which reflects the dynamics of alternative cost of money-keeping. The model for the Consumer Price Index also included the price index in the electric power industry, because the dynamics of manufacturers' costs relies heavily on this indicator.

The baseline indicator to be noted is the real exchange rate, which can influence the value of exports and imports, and its fluctuations can result in changes to the relative value of domestically-produced and imported goods, though the influence of this indicator turns out to be insignificant in econometric models. Global prices of exported resources, particularly crude oil prices, are most significant factors, which determine the dynamics of exports: a higher price leads to greater exports of goods. The level of personal income in the economy (labor costs) was used to describe the relative competitive power of Russian goods. Fictitious variables D12 and D01 – equal to one in December and January and zero in other periods – were added so that seasonal fluctuations were factored in. The dynamics of imports is effected by personal and corporate incomes whose increase triggers higher demand for all goods including imported ones. The real disposable cash income reflects the personal income; the Industrial Production Index reflects the corporate income.

The forecast values of foreign exchange rates were also calculated using structural models of their dependence on global crude oil prices.

The forecast values of explanatory variables, which are required for forecasting on the basis of structural models, were calculated using ARIMA models (p, d, q).

The paper also presents calculations of the values of the Industrial Production Index, the Producer Price Index and the Total Unemployment Index, which were calculated using the results of business surveys conducted by the Gaidar Institute. Empirical studies show¹ that the use of series of business surveys as explanatory variables² in forecasting models can make forecasting more accurate on the average. Future values of these indicators were calculated using ADL-models (seasonal autoregressive delays were added).

The Consumer Price Index and the Producer Price Index are also forecast using large datasets (factor models – FM). The construction of factor models relies basically on the evaluation of the principal components of a large dataset of socio-economic indicators (112 indicators in this case). The lags of these principal components and the lags of the explanatory variable are used as explanatory variables in these models. A quality analysis of the forecasts obtained for different configurations of the factor models was used to chose a model for the CPI, which included 9th, 12th and 13th lags of the four principal components, as well as 1st and 12th lags of the variable itself, and a model for the PPI, which included 8th, 9th and 12th lags of the four principal components, as well as 1st, 3rd and 12th lags of the variable itself.

All calculations were performed using the Eviews econometric package.

¹ See, for example: V. Nosko, A. Buzaev, P. Kadochnikov, S. Ponomarenko. The Analysis of Forecasting Parameters of Structural Models and Models with Business Surveys' Findings. Moscow, IEP, 2003.

² Used as explanatory variables were the following series of the business surveys: the current/expected change in production, the expected changes in the solvent demand, the current/expected price changes and the expected change in employment.

Table 1

CALCULATIONS OF FORECAST VALUES OF INDICES OF INDUSTRIAL PRODUCTION¹, (%)

	for manufac- ies (electricity, water, and gas) IIP for utilit- ies (electricity, products petroleum fabricated metal machinery products	HAGE Rosstat HSE R	respective month of the previous year	3 1.8 0.8 0.7 2.5 1.7 -3.0 -0.9 -0.4 -3.9 -1.1 -0.2	3.0 4.0 6.9 1.6 3.0 -4.0 -3.3 1.7 -2.6 -7.1 5.4 5.4	. 0.7 -2.5 -0.9 3.4 3.1 -1.1 0.0 -1.1 1.4 9.1 18.8	.1.3 0.8 8.0 2.8 1.6 2.1 -1.8 -2.1 -1.6 15.3 18.7	i 0.4 2.5 6.2 2.4 2.0 -0.6 0.0 0.9 0.7 9.7 0.8	1.3 4.8 4.0 2.5 2.6 3.1 -1.2 3.1 2.3 20.9 0.3	growth in 2015/2016 on the respective month of 2014/2015	3 -3.3 -3.5 -2.7 2.4 1.3 -1.8 -3.8 -6.5 -4.0 -5.8 9.5	1 -4.2 -6.3 -6.4 3.1 1.0 1.9 -0.6 -13.3 -6.5 -10.9 -11.7	3 -4.5 2.5 4.6 2.0 3.0 -3.2 -6.7 -6.5 -7.3 2.2 -10.3	0.0 0.0 0.0 4.5 4.9 -1.9 -4.1 -1.4 0.4 1.0 -9.5	2.0 80. 40.	1:01 0:0- 1:0- 0:0- 1:0-
		<u> </u>	year							of 2014/20					3.7 -9.	
	IIP for fo product		he previous	2.5						sctive month					2.5	
	or utilit- ectricity, and gas)	NRU	month of t	0.7	6.9	-0.9	8.0	6.2	4.0	in the respe	-2.7	-6.4	4.6	-0.8	9.0-	
-	IIP fo ies (ele water,	Rosstat	espective	0.8	4.0	-2.5	8.0	2.5	4.8	15/2016 0	-3.5	-6.3	2.5	0.0	-0.8	
1101	nanufac- ing		on the re	1.8	2.0	0.7	-1.3	0.4	1.3	vth in 20.	-3.3	-4.2	-4.5	0.0	-0.4	
	IIP for r tur	Rosstat	ed growth	-0.3	9.0	0.3	0.2	1.8	2.3		-5.3	-6.1	-5.6	-1.0	-2.8	
-	IIP for mining	NRU	Expect	1.5	2.2	6.0	-2.7	-2.3	0.3	For reference: actual	1.7	1.3	8.0	6.4	4.9	
	IIP for	Rosstat		9.0	-0.2	0.3	-2.2	-1.5	6.0	For refe	-0.1	0.1	0.4	5.8	4.2	
) i	uction	NRU HSE		2.8	1.2	2.4	6.0	2.7	1.1		-1.7	-2.9	-1.0	1.9	1.3	
	Index of industrial production	NRU		2.0	3.0	2.1	1.0	1.3	2.1		-1	-2	-1	1	ij	
	of indust	Rosstat		1.4	0.2	1.8	-0.1	1.3	0.0		-3.5	-4.5	-2.7	1.0	-0.5	
	Index	Ros		0.1	-0.7	2.2	6.0	8.0	1.3		c.	7-	3-	1	Ÿ	
				Nov 16	Dec 16	Jan 17	Feb 17	Mar 17	Apr 17		Nov 15	Dec 15	Jan 16	Feb 16	Mar 16	

Note. In the time spans under review, the series of the Rosstat and the NRU HSE chain indices of IIP, as well as the NRU HSE chain IIP for manufacturing are identified as fabricated metal products, as well as the NRU HSE chain IIP for mining and Rosstat chain IIP for machinery and equipment are identified as stationary processes around the stationary processes around the trend with an endogenous structural change; the series of the Rosstat and the NRU HSE chain IIPs for manufacturing, for primary metals and trend with two endogenous structural changes. The time series of other chain indices are stationary at levels.

1 It is to be noted that for making of forecasts so-called "raw" indices (without seasonal and calendar adjustment) were used and for that reason in most models existence of the season factor is taken into account and, as a consequence, the obtained outputs reflect the seasonal dynamics of the series.

INDUSTRIAL PRODUCTION AND RETAIL SALES

Industrial production

For making forecast for November of 2016 – April of 2017, the series of monthly data of the indices of industrial production released by the Federal State Statistics Service (Rosstat) from January 2002 to August 2016, as well as the series of the base indices of industrial production released by the National Research University Higher School of Economics (NRU HSE¹) over the period from January 1999 to September 2016 were used (the value of January 2000 was equal to 100%). The forecast values of the series were calculated on the basis of ARIMA-class models. The forecast values of the Rosstat and the NRU HSE indices of industrial production are calculated using business surveys (BS) as well. The obtained results are shown in Table 1.

As seen from *Table 1*, the average² growth of the index of industrial production computed by the NRU HSE in November 2016 – April 2017 against the same period of the previous year amounts to 1.9% on industry as a whole. For the index of industrial production computed by Rosstat, this indicator constitutes 0.8%. As of end-2016, the forecast annual growth of the index of industrial production computed by Rosstat will amount to 0.1%, and the index of industrial production computed by the NRU HSE – 3.6%. The average monthly values of the index of industrial production for mining computed by Rosstat and the NRU HSE in November 2016 – April 2017 will come to (-0.4%) and 0.0%, respectively. In production of coke and petroleum products growth is forecast at (-0.6%) and (-1.2%) for Rosstat and the NRU HSE indices, respectively.

In November 2016 – April 2017 in comparison with the same period of last year, the average growth of the NRU HSE index of industrial production in manufacturing comes to 0.8% and the Rosstat index at 0.8%. The average monthly values of the Rosstat and the NRU HSE index for industrial production of food products constitute 2.5% and 2.3%, respectively. The average monthly values

of the index of industrial production for primary metals and fabricated metal products in November 2016 – April 2017 computed by Rosstat and the NRU HSE constitute 0.3% and (-0.7%), respectively. In manufacture of machinery and equipment, the average increase is forecast at 7.8% and 7.3% for the Rosstat and the NRU HSE indices, respectively.

The average growth of the index of industrial production for utilities (electricity, gas and water) computed by Rosstat for November 2016 – April 2017 in comparison with the same period of the previous year constitutes 1.7%; the same indicator for the NRU HSE index comes to 4.1%.

The Rosstat indices of industrial production across various types of economic activity in 2016 will average (across types of activity) 0.5%, the NRU HSE indices of industrial production will grow by 1.9%.

Retail Sales

This section (Table 2) presents forecasts of monthly retail sales made on the basis of monthly Rosstat data over January 1999 – September 2016.

Table 2
CALCULATIONS OF FORECAST VALUES OF THE
RETAIL SALES AND THE REAL RETAIL SALES

KLIAIL SALLS AND THE KLAL KLIAIL SALLS											
For	recast value according to A	RIMA-model									
	Retail sales, billion RUR (in brackets – growth on the respective month of the previous year, %)	Real retail sales (as % of the respective period of the previous year)									
Nov 16	2490.3 (4.3)	96.8									
Dec 16	3148.6 (8.6)	97.2									
Jan 17	2152.8 (1.3)	95.5									
Feb 17	2122.3 (1.1)	95.2									
Mar 17	2317.2 (4.4)	95.8									
Apr 17	2294.5 (4.1)	95.5									
For re	eference: actual values in the	ne same months									
	of 2014/2015										
Nov 15	2387.3	87.8									
Dec 15	2898.1	85.9									
Jan 16	2126.2	93.6									
Feb 16	2098.6	95.3									
Mar 16	2220.3	93.8									
Apr 16	2204.2	94.9									

Note. The series of retail sales and real retail sales over January 1999 – September 2016.

¹ The indices in question are calculated by E.A. Baranov and V.A. Bessonov.

² The average growth of industrial production indices is understood here as the average value of the said indices for six forecast months.

CALCULATIONS OF FORECAST VALUES OF VOLUMES OF FOREIGN TRADE TURNOVER WITH COUNTRIES OUTSIDE THE CIS

	70	Percentage of actual data in the respective month of the previous year	$_{ m NM}$	66	86	69	103	26	96							
	Import from countries outside the CIS	Percentage of actual data in the respective month of the previous year	ARIMA	105	94	81	91	102	104		14.7	15.6	8.8	11.6	13.9	13.4
()	mport from	Forecast values (billion USD a month)	SM	14.6	15.3	6.1	12.0	13.4	12.8		14	13	8	11	15	15
		Forecast values (billion USD a month)	ARIMA	15.4	14.7	7.2	10.6	14.1	13.9							
		tage of ta in the e month	$_{ m SM}$	86	66	91	86	100	101	D)						
	countries the CIS	Percentage of actual data in the respective month of the previous year	ARIMA	84	84	80	91	105	94	billion US	ਲ	5.	∞.	£.	6.	18.6
	Export to countries outside the CIS	t values USD a	$_{ m SM}$	21.0	24.3	13.5	16.9	19.8	18.7	For reference: actual values in respective months of 2015/2016 (billion USD)	21.5	24.5	14.8	17.3	19.9	18
7 - 17 / 17		Forecast values (billion USD a month)	ARIMA	18.1	20.5	11.8	15.7	8.02	17.5	nonths of 2						
	sə	Percentage of actual data in the respective month of the previous year	$_{ m NM}$	86	92	79	124	86	96	espective n						
\	all countri		ARIMA	105	95	92	97	102	103	alues in re	ಸ್	4.	8	6.	4.	2.
	Import from all countries		$_{ m SM}$	16.1	16.0	7.8	16.0	15.1	14.6	se: actual	16.5	17.4	9.8	12.9	15.4	15.2
)	Im	Forecast values (billion USD a month)	ARIMA	17.2	16.5	9.1	12.5	15.8	15.6	or reference						
		tage of ta in the e month	$_{ m SM}$	102	66	66	97	97	105	F						
	Export to all countries	Percentage of actual data in the respective month of the previous year	ARIMA	83	82	98	92	101	66		4.	7.	1.	0.	.2	7.
() ()	Export to a	t values USD a nth)	$_{ m SM}$	26.0	28.6	16.9	19.5	22.4	22.8		25.4	28.7	17.1	20.0	23.2	21.7
	Ħ	Forecast values (billion USD a month)	ARIMA	21.1	23.5	14.6	18.5	23.4	21.5							
				Nov 16	Dec 16	Jan 17	Feb 17	Mar 17	Apr 17		Nov 15	Dec 15	Jan 16	Feb 16	Mar 16	Apr 16

Note. Over the period from January 1999 to September 2016, the series of export, import, export to the countries outside the CIS and import from the countries outside the CIS were identified as stationary series in the first-order differences. In all the cases, seasonal components were included in the specification of the models.

FORFIGN TRADE INDICES

As seen from $Table\ 2$, the average forecast growth of nominal volumes of monthly trade turnover amounts to around 4.0% for November 2016 – April 2017 against the corresponding period of 2015–2016.

The average forecast decrease of the monthly real trade turnover for the period from November 2016 – April 2017 against the same period of 2015-2016 constitutes 4.0%.

At an annual rate, forecast growth of the nominal index of retail trade turnover in 2016 will come to 8.6%, and in real terms will decrease by 5.0%.

FOREIGN TRADE INDICES

Model calculations of forecast values of the export and export to countries outside the CIS and the import and import from countries outside the CIS were made on the basis of the models of time series and structural models evaluated on the basis of the monthly data over the period from September 1998 to September 2016 on the basis of the data released by the Central Bank of Russia¹. The results of calculations are shown in Table 3.

The average forecast decrease of exports, imports, exports outside the CIS and imports from the countries outside the CIS for November 2016 – April 2017 against the same period of 2015–2016 will amount to 4.9%, 1.3%, 6.2%, and -3.8%, respectively. The average forecast surplus volume of the trade balance with all countries for 2016 will constitute \$88.0bn, which corresponds to a decrease by 40.8% on the same period of 2015.

DYNAMICS OF PRICES

The Consumer Price Index and Producer Price Index

This section presents calculations of forecast values of the consumer price index and producer price index (as regards both the industry in general and some types of its activities under the National Industry Classification Standard (NICS)) made on the basis of the time-series models evaluated on the basis of the data released by Rosstat over the period from January 1999 to August 2016². Table 4 presents the results of model calculations of forecast values over November 2016 and April 2017 in accordance with ARIMA models, structural models (SM) and models computed with the help of business surveys (BS).

The forecast average monthly growth of the consumer price index in November 2016 – April 2017 will come to 0.7%. The price growth of industrial goods for this period is forecast at an average monthly rate of 0.4%. Annual growth of the consumer price index on average across three models will come to 5.7%. The same indicator for the producer price index is forecast at 5.4%.

For the producer price indices of Russian National Classifier of Economic Activities (OKVED) computed for November 2016 – April 2017 the following average monthly growth rates are forecast: for mining and quarrying (-2.9%), manufacturing 0.5%, utilities (electricity, water, and gas) 0.8%, food products 0.9%, textile and sewing industry 0.6%, wood products 0.5%, pulp and paper industry 0.4%, coke and refined petroleum 1.3%, for chemical industry 0.0%, for basic metals and fabricated metal 0.2%, for machinery and equipment 0.9%, and for transport equipment and manufacturing 0.6%.

Annual growth of the producer price indices across types of economic activity will average 6.3%. By end-2016, maximum annual growth is forecast in the production of basic metals and fabricated metal (15.6%), and the minimum – in chemical production (-3.9%).

¹ The data on the foreign trade turnover is calculated by the CBR in accordance with the methods for making of the balance of payment in prices of the exporter-country (FOB) in billion USD.

² Structural models were evaluated in the period from October 1998.

CALCULATIONS OF FORECAST VALUES OF PRICE INDICES

	for transport equipment manufacturing		100.3	100.9	100.7	100.2	100.7	100.7		107.2	108.1	100.7	100.9	101.6	102.3		115.9	116.3	6.66	101.0	101.5	102.1
	for machinery and equipment		100.8	100.9	101.2	101.5	100.6	100.6		108.5	109.5	101.2	102.8	103.4	104.0		113.5	114.0	103.3	104.0	105.2	105.9
	or basic metals baricated fabricated metal		100.9	100.8	8.66	99.1	100.2	100.4		114.6	115.6	8.66	98.9	99.1	99.5		113.1	111.9	7.66	102.3	102.7	102.8
	for chemical industry		8.66	6.66	6.66	100.0	100.1	100.2		96.2	96.1	6.66	6.66	100.0	100.2		119.5	118.2	100.2	101.9	101.5	100.4
	for coke and refined petroleum		101.2	101.5	101.4	100.9	101.2	101.5		106.2	107.8	101.4	102.3	103.5	105.0	()	108.7	100.9	92.6	92.7	97.2	9.96
	bns qluq rot yytsubni rəqaq		6.66	100.5	100.6	100.6	100.5	100.5		107.8	108.3	100.6	101.2	101.7	102.2	2014/2015	118.2	119.0	101.2	106.1	106.9	107.2
dices:	boow rot stouborq		100.4	100.3	100.6	100.5	100.8	100.7		104.7	105.0	100.6	101.1	101.9	102.6	of 2015/2016 (% of December 2014/2015)	109.3	109.3	100.5	100.2	101.3	101.7
Producer price indices:	for textile and stry	nonth)	100.6	100.2	100.7	100.6	100.6	100.7	5/2016)	108.0	108.2	100.7	101.4	102.0	102.7	16 (% of D	113.6	113.6	102.0	103.3	104.6	104.9
Produce	boof rof stoubord	revious m	100.4	100.7	100.7	101.1	101.2	101.1	nber 2015	105.8	106.5	100.7	101.8	103.0	104.1	2015/201	113.2	113.9	100.8	101.3	101.4	101.8
	for utilities (electricity, water, and gas)	Forecast values (% of the previous month)	100.3	100.1	100.9	102.7	100.9	6.66	(% of December 2015/2016)	104.4	104.4	100.9	103.7	104.6	104.5	periods of	109.2	109.4	99.3	100.5	100.3	6.66
	rof gnirutəslunsm	t values (100.4	100.2	101.0	100.5	100.3	100.5		107.9	108.1	101.0	101.5	101.8	102.3	s in the same periods	112.7	111.2	99.4	100.6	101.9	102.0
	bns gninim rot gniyrrsup	Forecast	92.6	98.4	98.5	91.5	102.0	96.3	Forecast values	99.1	97.5	98.5	90.1	91.9	88.6		118.0	109.8	92.6	83.8	94.3	106.6
	lsirtsubni rot (MT) sboog		100.8	8.66	100	101.4	101.8	101.4		105.6	105.4	100.0	101.4	103.2	104.7	For reference: actual value						
	lsirtsubni rot (SA) sboog		7.66	99.4	8.66	6.66	100.5	100.3		106.2	105.6	8.66	2.66	100.2	100.5	reference	113.2	110.8	98.6	97.0	100.0	102.6
	lairtsubni rot (AMIAA) sboog		99.4	100.1	100.0	100.2	101.6	100.4		105.1	105.2	100.0	100.2	101.8	102.2	For						
	The consumer MA) xəbni əsirq		100.6	100.6	101.7	100.7	100.8	100.9		105.0	105.6	101.7	102.4	103.2	104.2							
	The consumer AS) xəbni əsirq		100.5	100.6	100.5	100.4	100.5	100.4		105.0	105.6	100.5	100.9	101.4	101.8		112.3	113.2	101.0	101.6	102.1	102.5
	The consumer price index (AMIAA)		100.5	100.4	101.3	100.6	100.6	100.7		105.3	105.8	101.3	101.9	102.6	103.3							
			Nov 16	Dec 16	Jan 17	Feb 17	Mar 17	Apr 17		Nov 16	Dec 16	Jan 17	Feb 17	Mar 17	Apr 17		Nov 15	Dec 15	Jan 16	Feb 16	Mar 16	Apr 16

Note. Over the period from January 1999 to August 2016, the series of the chain producer price index for machinery are identified as a stationary process around the trend with two endogenous structural changes. The series of other chain price indices are stationary at levels.

The Cost of the Monthly per Capita Minimum Food Basket

This section presents calculations of forecast values of the cost of the monthly per capita minimum food basket over November 2016 and April 2017. The forecasts were made based on time series with use the Rosstat data over the period from January 2000 to September 2016. The results are shown in Table 5.

As can be seen from *Table 5*, cost growth of the minimum set of food products is forecast compared to the corresponding period of the previous year. Herewith, forecast cost of the minimum set of food products constitutes around RUR 3,701.7. The forecast cost growth of the minimum set of food products will average around 2.1% compared to the level of the corresponding period of the previous year. Annual growth of the cost of the minimum set of food products in 2016 will constitute 1.8%.

Indices of Freight Rates

This section presents calculations of forecast values of freight rate indices on cargo carriage¹, made on the basis of time-series models evaluated on the Rosstat data over the period from September 1998 to August 2016. Table 6 shows the results of

Table 5
THE FORECAST OF THE COST OF THE MONTHLY
PER CAPITA MINIMUM FOOD BASKET

Forecast values acco	ording to ARIMA-model (RUR)
Nov 16	3620.0
Dec 16	3655.0
Jan 17	3689.2
Feb 17	3721.2
Mar 17	3746.7
Apr 17	3778.0
	al values in the same months
of 2015/	2016 (billion RUR)
Nov 15	3547.2
Dec 15	3589.9
Jan 16	3627.1
Feb 16	3649.8
Mar 16	3655.3
Apr 16	3677.6
Expected growt	th on the respective month
of the	previous year (%)
Nov 16	2.1
Dec 16	1.8
Jan 17	1.7
Feb 17	2.0
Mar 17	2.5
Apr 17	2.7

Note. The series of the cost of the monthly per capita minimum food basket over the period from January 2000 to September 2016 are stationary in the first-order differences.

model calculations of forecast values in November 2016 – April 2017. It should be noted that some of the indices under review (for instance, the pipeline rate index) are adjustable ones and for that reason their behavior is hard to describe by means of the time-series models. As a result, the future values may differ greatly from the real ones in case of the centralized increase of rates in the period of forecasting or in case of absence of such an increase in the forecasting period, but with it taking place shortly before the beginning of that period.

CALCULATIONS OF FORECAST VALUES OF INDICES OF FREIGHT RATES

The composite The index The index freight rate index of truckload freight rate of pipeline rate Forecast values according to ARIMA-models (% of the previous month) Nov 16 100.1 100.0 100.6 Dec 16 100.1 100.0 102.3 Jan 17 100.1 101.7 100.9 Feb 17 100.0 100.1 99.8 Mar 17 100.0 99.9 101.5 Apr 17 103.8 99.9 101.8 Forecast values according to ARIMA-models (% of December of the previous year) Nov 16 101.4 119.3

Table 6

¹ The paper presents a review of the composite freight rate index on freight transport and the truckload freight rate index, as well as the pipeline rate index. The composite freight rate index is computed on the basis of the freight rate indices by individual types of transport: rail, pipeline, shipping, domestic water-borne, and truckload freight and air service (for more detailed information, pls. refer, for instance, to: *Prices in Russia*. The Official Publication of Goskomstat of RF, 1998).

Table 6, cont'd

	The composite	The index	The index
	freight rate index	of truckload freight rate	of pipeline rate
Dec 16	111.7	101.4	122.0
Jan 17	100.1	101.7	100.9
Feb 17	100.2	101.7	100.7
Mar 17	100.2	101.6	102.2
Apr 17	104.0	101.5	104.0
	For reference: actual values i	n the same period of $2015/2016$ (% of	the previous month)
Nov 15	100.2	100.2	100.2
Dec 15	100.6	101.9	100.5
Jan 16	100.7	102.2	93.7
Feb 16	99.8	100.1	99.8
Mar 16	99.5	99.0	99.6
Apr 16	108.9	99.6	119.2

Note. Over the period from September 1998 to August 2016, the series of the freight rates index were identified as stationary ones; the other series were identified as stationary ones over the period from September 1998 to August 2016, too; fictitious variables for taking into account particularly dramatic fluctuations were used in respect of all the series.

According to the forecast results for November 2016 – April 2017, the composite freight rate index will increase on average 0.7% per month. In April 2017, seasonal growth of the index is expected by 3.8 p.p. Its annual growth in 2016 will come to 11.7%.

The index of truckload freight rate will grow at a monthly average rate of 0.3% in the course of given six months. Its annual growth is forecast at 1.4% in 2016.

Pipeline transport index will be growing in the course of next six months at a monthly average rate of 1.2%. As a result, its annual growth will amount to 22.0% in 2016. In April 2017, seasonal growth of 1.8 p.p. is expected.

World Prices of Natural Resources

This section presents calculations of such average monthly values of Brent crude prices (US\$ per barrel), the aluminum prices (US\$ per ton), the gold prices (\$ per ounce), the copper prices (US\$ per ton) and the nickel prices (US\$ per ton) over November 2016 and April 2017 as were received on the basis of nonlinear models of time series evaluated on the basis of the IMF data over the period from January 1980 to September 2016.

Table 7
CALCULATIONS OF FORECAST VALUES OF WORLD PRICES ON NATURAL RESOURCES

	ON LOOP MOTOR OF TORIEST WILL OF THE RESTORAGE											
	Brent oil	Aluminum	Gold	Copper	Nickel							
	(\$ per barrel)	(\$ per ton)	(\$ per ounce)	(\$ per ton)	(\$ per ton)							
		F	'orecast values									
Nov 16	45.70	1583	1353	4737	10470							
Dec 16	44.89	1552	1369	4745	10461							
Jan 17	45.20	1548	1369	4757	10608							
Feb 17	44.97	1545	1374	4770	10616							
Mar 17	44.89	1530	1393	4786	10647							
Apr 17	45.23	1530	1410	4804	10652							
	Exp	ected growth on the re	spective month of the	previous year (%)								
Nov 15	2.9	7.8	24.7	-1.3	13.3							
Dec 15	19.0	3.7	28.2	2.3	20.1							
Jan 16	46.8	4.5	24.8	6.4	24.7							
Feb 16	35.5	0.9	14.5	3.7	27.9							
Mar 16	14.9	-0.1	11.8	-3.4	22.1							
Apr 16	7.1	-2.6	13.5	-1.4	20.0							
	F	or reference: actual va	lues in the same perio	od of 2015/2016								
Nov 15	44.42	1468	1086	4800	9244							

Table 7, cont'd

	Brent oil (\$ per barrel)	Aluminum (\$ per ton)	Gold (\$ per ounce)	Copper (\$ per ton)	Nickel (\$ per ton)
Dec 15	37.72	1497	1068	4639	8708
Jan 16	30.8	1481	1097	4472	8507
Feb 16	33.2	1531	1200	4599	8299
Mar 16	39.07	1531	1246	4954	8717
Apr 16	42.25	1571	1242	4873	8879

Note. Over the period from January 1980 to September 2016, the series of prices of crude oil, nickel, gold, copper and aluminum are series of DS type.

The average forecast of crude oil price amounts to around \$45.1 per barrel, which is above its corresponding year-earlier indices on average by 21.0%. Aluminum prices are forecast at around \$1,548.0 per ton and their average forecast growth constitutes around 2.0% compared to the same level last year. Forecast for gold prices constitute around \$1,378.0 per ounce. Forecast average copper prices constitute around \$4,766 per ton and of nickel prices – around \$10,575 per ton. Average forecast price growth on gold constitutes around 20.0%, average reduction of copper prices – around 1.0%, average reduction of nickel prices – 21% compared to the corresponding level of last year.

By end-2016, forecast growth of prices on crude oil, aluminum, gold, copper and nickel against end-2015 according to the forecast will come to 19%, 3.7%, 28.2%, 2.3%, and 20.1% respectively.

MONETARY INDICES

The future values of the monetary base (in the narrow definition: cash funds and the Fund of Mandatory Reserves (FMR) and $M_{\rm 2}$ monetary aggregate over the period from November 2016 to

April 2017 were received on the basis of models of time-series of respective indices calculated by the CBR¹ over the period from October 1998 to October (September – for M2 time series) 2016. Table 8 presents the results of calculations of forecast values and actual values of those indices in the same period of previous year. It is to be noted that due to the fact that the monetary base is an instrument of the CBR policy, forecasts of the monetary base on the basis of time-series models are to a certain extent notional as the future value of that index is determined to a great extent by decisions of the CBR, rather than the inherent specifics of the series.

In November 2016 – April 2017, the monetary base will be growing at an average monthly rate of 0.8%, and the monetary indicator M2 – at an average monthly rate of 0.9%. In January 2017, seasonal growth of monetary base if forecast by 5.8% and monetary indicator M2 – by 2.9%.

In 2016, annual growth of the indicator $\rm M_{_2}$ is forecast at the level of 12.5%, and the monetary base – 9.2%.

Table 8 THE FORECAST OF $\rm M_{\rm 2}$ AND THE MONETARY BASE

7 (178) 1112 1710 17217 (171 87 102												
	The I	Monetary base		M_{2}								
	Billion RUR	Growth on the previous month, %	Billion RUR	Growth on the previous month, %								
Nov 16	8710	1.8	37154	0.4								
Dec 16	8675	-0.4	37477	0.9								
Jan 17	9177	5.8	38549	2.9								
Feb 17	8803	-4.1	38714 0.4									
Mar 17	8965	1.8	38879	0.4								
Apr 17	8933	-0.4	39044	0.4								
		actual value in (growth on the										
Nov 15		0.2		-0.3								
Dec 15		-0.1		1.4								
Jan 16		10.1		7.5								
Feb 16		-6.3	-2.7									
Mar 16		1.0	0.8									
Apr 16		-0.7	1.0									

Note. Over the period from October 1998 to October (September) 2016, all the time series of monetary indices were attributed to the class of series which are stationary in the first-order differences and have an explicit seasonal component.

¹ The data on the specific month is given in accordance with the methods of the CBR as of the beginning of the following month.

INTERNATIONAL RESERVES

This section presents the outputs of the statistical estimation of such future values of the international reserves of the Russian Federation¹ as were received on the basis of evaluation of the model of time series of the gold and foreign exchange reserves on the basis of the data released by the CBR over the period from October 1998 to October 2016. That index is forecast without taking into account a decrease in the amount of reserves due to foreign debt payment and for that reason the values of the volumes of the international reserves in the months where foreign debt payments are made may happen to be overestimated (or, otherwise, underestimated) as compared to the actual ones.

Subsequent to the forecast results in November 2016 – April 2017, the international reserves will be growing by an average monthly rate of 0.4%. In 2016, annual growth of the international reserves will come to 9.6%.

Table 9
THE FORECAST OF THE INTERNATIONAL
RESERVES OF THE RUSSIAN FEDERATION

	Forecast values according to ARIMA-model							
Period	Billion USD	Growth on the previous month, %						
Nov 16	398.6	0.2						
Dec 16	399.8	0.3						
Jan 17	401.7	0.5						
Feb 17	403.5	0.5						
Mar 17	405.3	0.4						
Apr 17	407.0	0.4						
For r	eference: actua	al values in the same period						
	of	2015/2016						
Nov 15	369.6	-0.4						
Dec 15	364.7	-1.3						
Jan 16	368.4	1.0						
Feb 16	371.6	0.9						
Mar 16	380.5	2.4						
Apr 16	387.0	1.7						

Note. Over the period from October 1998 to October 2016, the series of the gold and foreign exchange reserves of the Russian Federation were identified as stationary series in difference.

FOREIGN EXCHANGE RATES

The model calculations of prospective values of the foreign exchange rates (RUR per USD and USD per euro) were made on the basis of assessment of the time series models (ARIMA) and structural models (SM) of the relevant indicators released by the Central Bank of Russia as of the last date of each month over the periods from October 1998 to October 2016 and from January 1999 to October 2016², respectively.

USD/RUR average exchange rate in the intervening period is forecast on average along two models in the amount of RUR 63.97 for USD. Forecast by end-2016 average (along two models) indicator will total Rb 63.08 for USD.

Euro/USD average exchange rate is forecast at USD 1.10 per 1 euro. By end-2016, the indicator is forecast at USD 1.10 per 1 euro along two models.

FORECASTS OF THE USD/RUR AND EUR/USD EXCHANGE RATES

	The US	SD/RUR	The EUR/USD					
Period		ige rate	exchange rate					
	(RUR p	er USD)	(USD per EUR)					
	ARIMA	SM	ARIMA	SM				
Nov 16	63.08	63.07	1.10	1.11				
Dec 16	63.63	63.62	1.10	1.10				
Jan 17	63.85	63.83	1.10	1.11				
Feb 17	64.13	64.18	1.10	1.11				
Mar 17	64.39	64.48	1.10	1.11				
Apr 17	64.66	64.74	1.10	1.11				
For refe	erence: actu	ıal values in	the similar	r period of				
		2015/2016	3					
Nov 15	66	.24	1.	05				
Dec 15	72	.88	1.	09				
Jan 16	75	.17	1.09					
Feb 16	75	.09	1.09					
Mar 16	67	.61	1.14					
Apr 16	64	.33	1.14					

Note. Over the respective periods, the series under review were identified as integrated series of the first order with a seasonal component.

¹ The data on the volume of the gold and foreign exchange reserves is presented as of the first day of the following month

² The authors use the IMF data over the period from January 1999 to August 2016. The data over the period from September and October 2016 was obtained from the foreign exchange rate statistics website: www.oanda.com

THE LIVING STANDARD INDICES

This section (Table 12) presents calculations of forecast values of indices of real wages, real disposable income and real income as were received on the basis of the model of time series of respective indices computed by Rosstat and taken over the period from January 1999 to September 2016. The above indices depend to a certain extent on the centralized decisions on raising of wages and salaries to public sector workers, as well as those on raising of pensions, scholarships and allowances; such a situation introduces some changes in the dynamics of the indices under review. As a result, the future values of the indices of real wages and real disposable income calculated on the basis of the series which last observations are either considerably higher or lower than the previous ones due to such a raising may differ greatly from those which are implemented in reality.

According to the results presented in *Table 11*, the expected average monthly fall of the real disposable cash income will constitute 1.9% per month compared to the same period of last year; the real cash income -1.9%, growth of the real wage is forecast at the average level of 0.9%.

By end-2016, forecast decrease of the real disposable cash income will amount to 4.9%; the real cash income – by 4.7%, and growth of the level of the real wages – 1.1%.

THE FORECAST OF THE LIVING STANDARD INDICES

Table 11

	Real disposable cash income	Real cash income	Real accrued wages
	Forecast values according to ARIMA	A-models (% of the respective r	month of 2015/2016)
Nov 16	97.5	97.6	104.4
Dec 16	96.4	96.3	103.8
Jan 17	99.2	98.7	103.3
Feb 17	98.0	98.3	100.3
Mar 17	97.6	98.0	102.0
Apr 17	100.0	99.9	101.6
For refe	rence: actual values in the respectiv	ve period of 2015/2016 (% of th	e same period of 2014/2015)
Nov 15	93.7	93.5	89.6
Dec 15	99.1	98.5	91.6
Jan 16	94.3	95.1	96.4
Feb 16	95.7	95.3	100.6
Mar 16	98.7	98.0	101.5
Apr 16	93.0	93.4	98.9

Note. For calculating purposes, the series of the real disposable cash income, real cash income and real accrued wages in the base form were used (January 1999 was adopted as a base period). Over the period from January 1999 to September 2016 those series were attributed to the class of processes which are stationary in differences and have an explicit seasonal component.

EMPLOYMENT AND UNEMPLOYMENT

For the purpose of calculation of the future values of the employment (of the number the gainfully employed population) and the unemployment (the total number of the unemployed), models of the time series evaluated over the period from October 1998 to August 2016 on the basis of the monthly data released by Rosstat² were used. The unemployment was calculated on the basis of the models with results of the findings from business surveys³, too.

¹ Real cash income is a relative index which is calculated by means of division of the index of the nominal size (which was actually formed in the period under review) of households' cash income by the CPI. Real disposable cash income is cash income minus mandatory payments and contributions. (See: Rossiisky Statistichesky Ezhegodnik, Moscow, Rosstat, 2004, p. 212).

² The index is computed in accordance with the methods of the International Labor Organization (ILO) and is given as of the month-end.

³ The model is evaluated over the period from January 1999 to August 2016.

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It is to be noted that feasible logical inconsistencies¹ in forecasts of employment and unemployment which totals should be equal to the index of economically active population may arise due to the fact that each series is forecast individually and not as a difference between the forecast values of the economically active population and another index.

Table 12 CALCULATION OF FORECAST VALUES OF THE INDICES THE EMPLOYMENT AND THE UNEMPLOYMENT

	Emplo	yment (ARIMA)		Unemployment	(ARIMA)		Unemploymen	nt (BS)
	Million people	Growth on the respective month of previ- ous year (%)	Million people	Growth on the respective month of pre- vious year (%)	% of the index of the number of the gainfully employed popu- lation	Million people	Growth on the respective month of previ- ous year (%)	% of the index of the number of the gainfully employed popu- lation
Nov 16	72.6	0.6	4.2	-4.4	5.8	4.2	-3.7	5.8
Dec 16	72.4	0.2	4.3	-2.3	5.9	4.3	-1.9	5.9
Jan 17	72.0	1.0	4.5	1.5	6.2	4.3	-1.9	6.0
Feb 17	72.0	0.7	4.5	3.0	6.3	4.4	-0.6	6.1
Mar 17	72.2	0.8	4.5	-3.0	6.2	4.4	-4.3	6.1
Apr 17	72.4	0.9	4.3	-4.1	6.0	4.4	-3.1	6.1
		For reference	: actual	values in the sar	me periods of 201	5/2016 (r	nillion people)	
Nov 15		72.2			4	1.4		
Dec 15		72.3			4	1.4		
Jan 16		71.3				1.4		
Feb 16		71.5			4	1.4		
Mar 16		71.6			4	4.6		
Apr 16		71.8			4	4.5		

Note. Over the period from October 1998 to August 2016, the series of employment is a stochastic process which is stationary around the trend. The series of unemployment is a stochastic process with the first order integration. Both indices include seasonal component.

According to ARIMA-model forecast (*Table 12*), in November 2016 – April 2017, growth of the number of employed in the economy on average will constitute 0.7% per month against the corresponding period of the previous year. Forecast by end-2016 indicator of the number of employed in the economy constitutes 72.4 mn persons.

The average fall of the total number of jobless is forecast at 2.1% per month against the corresponding period last year. Average number of jobless by end-2016 is forecast at 4.3 mn persons.

¹ For example, deemed as such a difference may be a simultaneous decrease both in the employment and the unemployment. However, it is to be noted that in principle such a situation is possible provided that there is a simultaneous decrease in the number of the economically active population.

ANNEX

Diagrams of the Time Series of the Economic Indices of the Russian Federation

Fig. 1a. The Rosstat industrial production index (ARIMA-model) (% of December 2001)

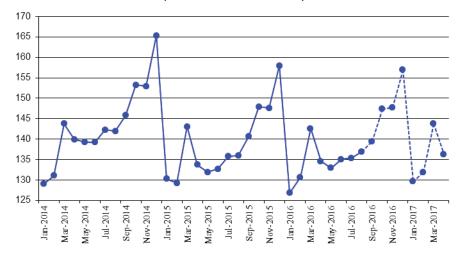


Fig. 1b. The NRU HSE industrial production index (ARIMA-model) (% of January 2005)

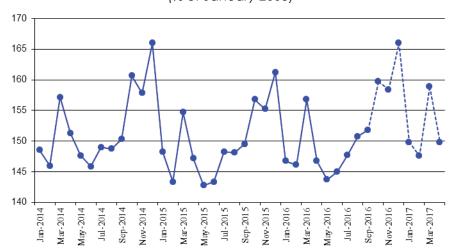


Fig. 2a. The Rosstat industrial production index for mining (% of December 2001)

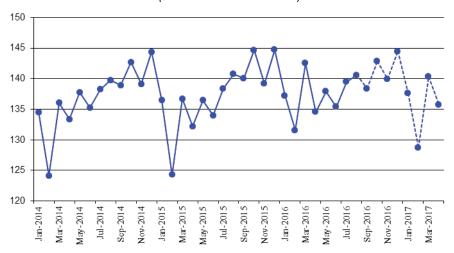


Fig. 2b. The NRU HSE industrial production index for mining (% of January 2005)

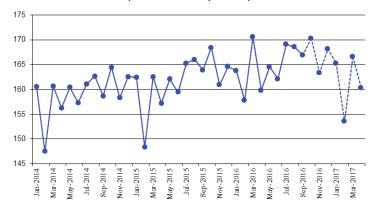


Fig. 3a. The Rosstat industrial production index for manufacturing (% of December 2001)

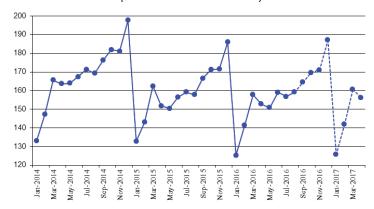


Fig. 3b. The NRU HSE industrial production index for manufacturing (% of January 2005)



Fig. 4a. The Rosstat industrial production index for utilities (electricity, water, and gas) (as a percentage of that in December 2001)

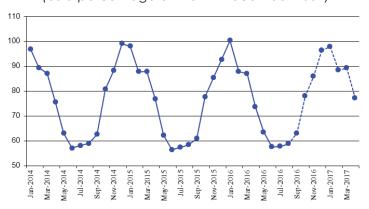


Fig. 4b. The NRU HSE industrial production index for utilities (electricity, water, and gas) (as a percentage of that in January 2005)

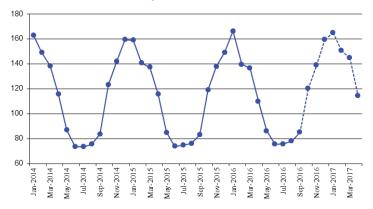


Fig. 5a. The Rosstat industrial production index for food products (as a percentage of that in December 2001)

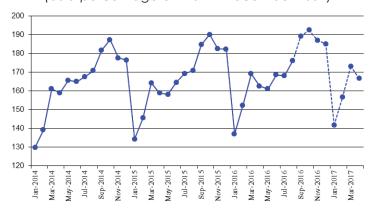


Fig. 5b. The NRU HSE industrial production index for food products (as a percentage of that in January 2005)

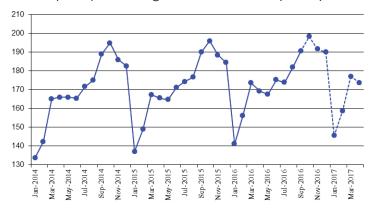
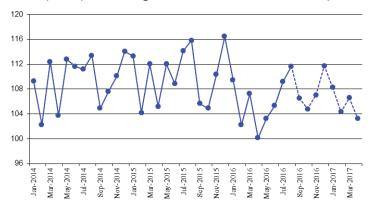


Fig. 6a. The Rosstat industrial production index for coke and petroleum (as a percentage of that in December 2001)



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Fig. 6b. The NRU HSE industrial production index for petroleum and coke (as a percentage of that in January 2005)

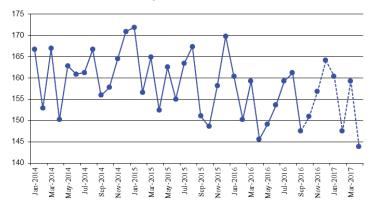


Fig.7a. The Rosstat industrial production index for primary metals and fabricated metal products (as a percentage of that in December 2001)

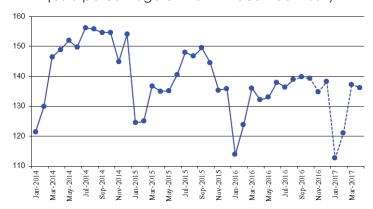


Fig. 7b. The NRU HSE industrial production index for primary metals and fabricated metal products (as a percentage of that in January 2005)

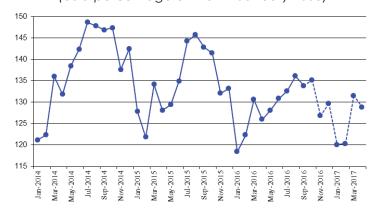


Fig. 8a. The Rosstat industrial production index for machinery (as a percentage of that in December 2001)

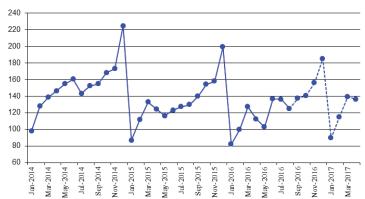


Fig. 8b. The NRU HSE industrial production index for machinery (as a percentage of that in January 2005)

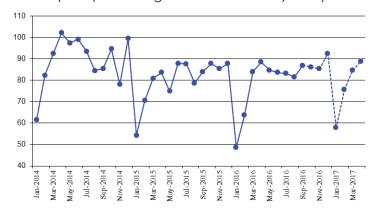


Fig. 9. The volume of retail sales (billion RUR)

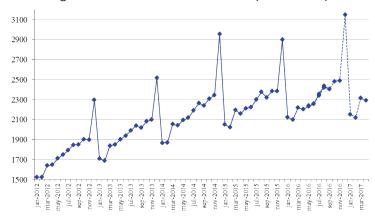


Fig. 9a. The real volume of retail sales (as a percentage of that in the same period of the previous year)

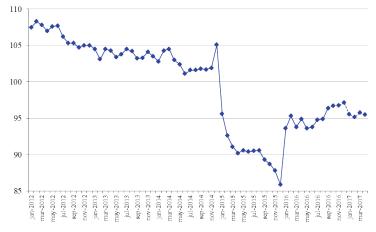


Fig. 10. Export to all countries (billion USD)

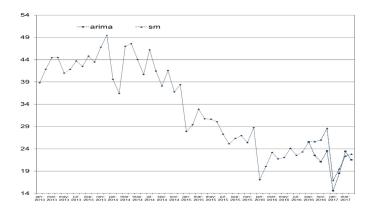


Fig. 11. Export to countries outside the CIS (billion USD)

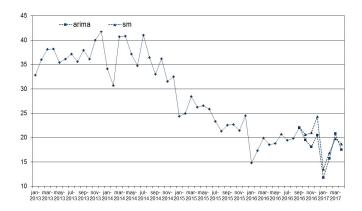


Fig. 12. Import from all countries (billion USD)

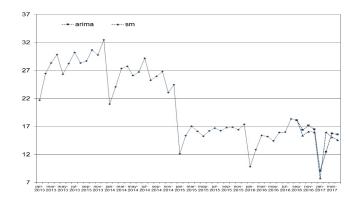


Fig. 13. Import from countries outside the CIS (billion USD)

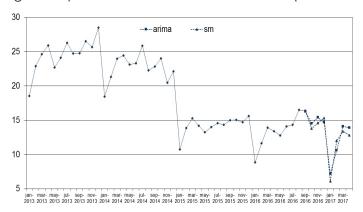


Fig. 14. The consumer price index (as a percentage of that in December of the previous year)

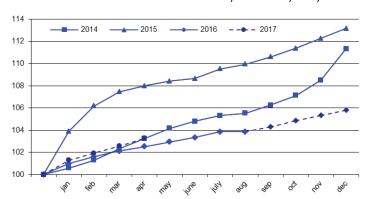


Fig. 14a. The consumer price index (as a percentage of that in December of the previous year) (SM)

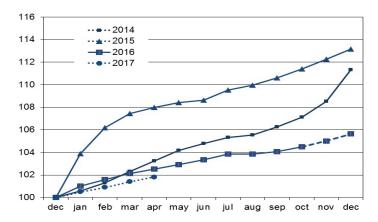


Fig. 15. The producer price index for industrial goods (as a percentage of that in December of the previous year)

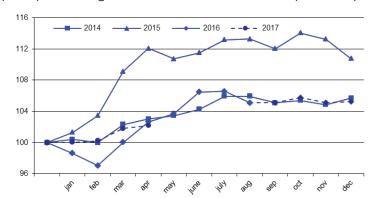


Fig. 16. The price index for mining (as a percentage of that in December of the previous year)

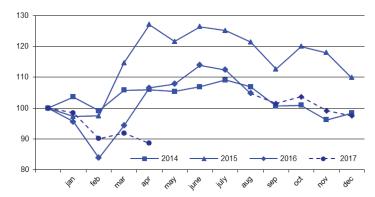


Fig. 17. The price index for manufacturing (as a percentage of that in December of the previous year)

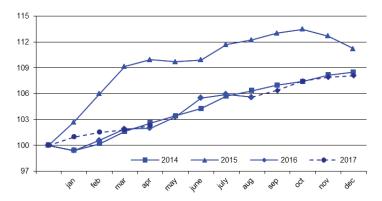


Fig. 18. The price index for utilities (electricity, water, and gas) (as a percentage of that in December of the previous year)

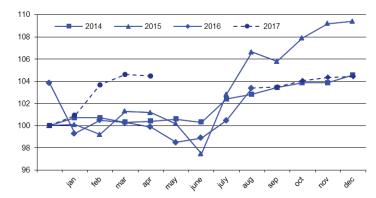


Fig. 19. The price index for food products (as a percentage of that in December of the previous year)

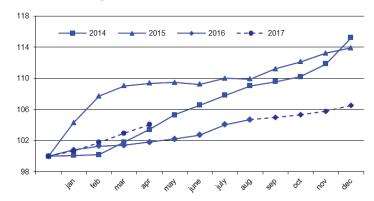


Fig. 20. The price index for the textile and sewing industry (as a percentage of that in December of the previous year)

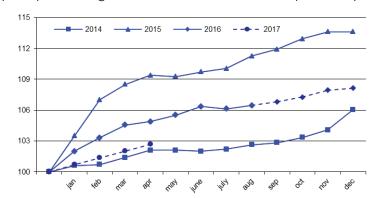


Fig. 21. The price index for wood products (as a percentage of that in December of the previous year)

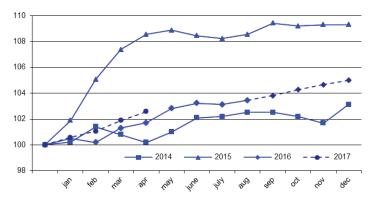


Fig. 22. The price index for the pulp and paper industry (as a percentage of that in December of the previous year)

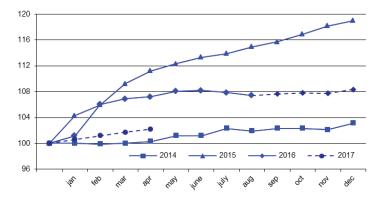


Fig. 23. The price index for coke and petroleum (as a percentage of that in December of the previous year)

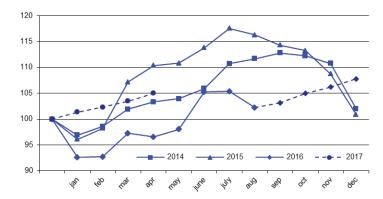


Fig. 24. The price index for the chemical industry (as a percentage of that in December of the previous year)

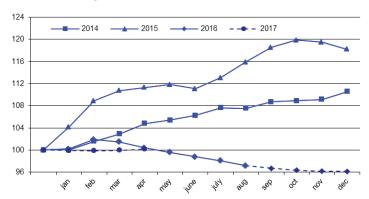


Fig.25. The price index for primary metals and fabricated metal products (as a percentage of that in December of the previous year)

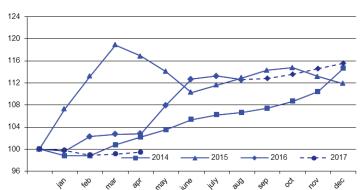


Fig.26. The price index for machinery (as a percentage of that in December of the previous year)

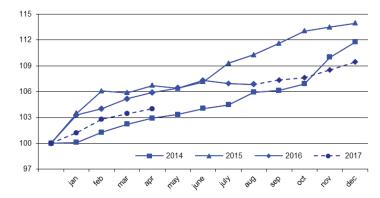


Fig.27. The price index for transport equipment manufacturing (as a percentage of that in December of the previous year)

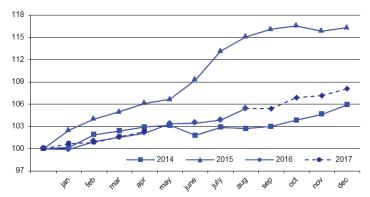


Fig. 28. The cost of the monthly per capita minimum food basket (RUR)

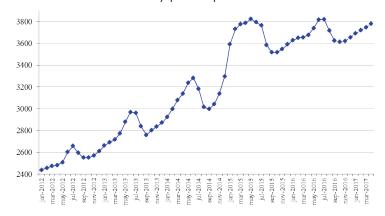


Fig. 29. The composite index of transport tariffs (for each year, as a percentage of that in the previous month)

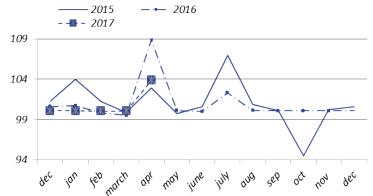


Fig. 30. The index of motor freight tariffs (for each year, as a percentage of that in the previous month)

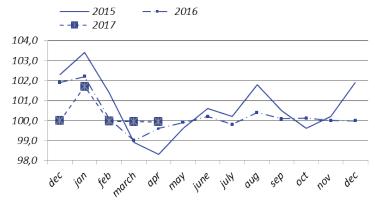


Fig. 31. The index of pipeline tariffs (for each year, as a percentage of that in the previous month)

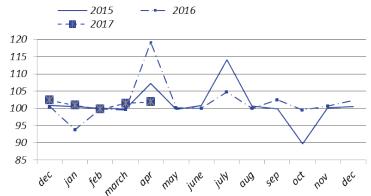


Fig. 32. The Brent oil price (\$ per barrel)

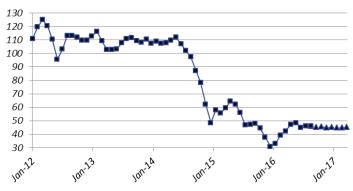


Fig. 33. The aluminum price (\$ per ton)

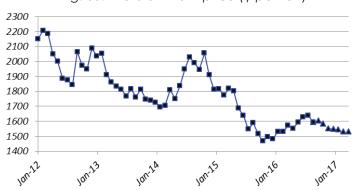




Fig. 35. The nickel price (\$ per ton)

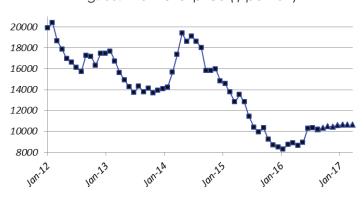


Fig. 36. The copper price (\$ per ton)

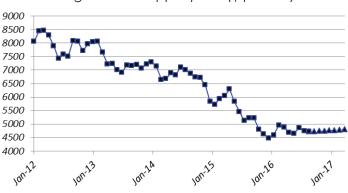
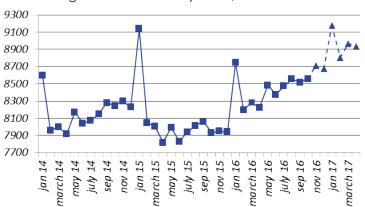


Fig. 37. The monetary base, billion RUR



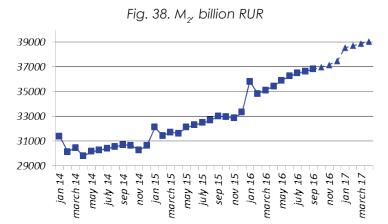


Fig. 39. The international reserves of the Russian Federation, million USD

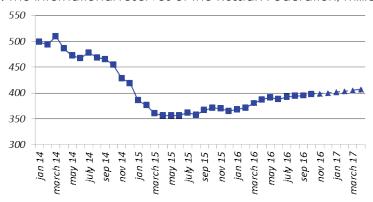


Fig. 40. The RUR/USD exchange rate

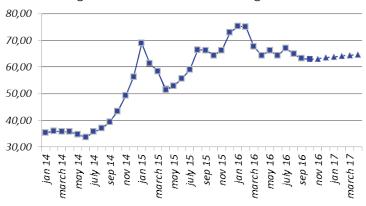
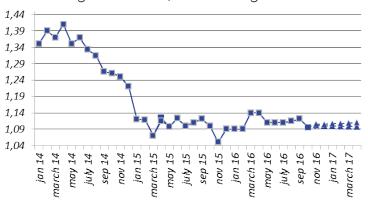


Fig. 41. The USD/EUR exchange rate



10'2016 MODEL CALCULATIONS OF SHORT-TERM FORECASTS

Fig. 42. Real disposable cash income (as a percentage of that in the same period of the previous year)



Fig. 43. Real cash income (as a percentage of that in the same period of the previous year)

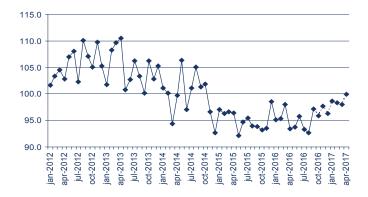


Fig. 44. Real accrued wages (as a percentage of those in the same period of the previous year)

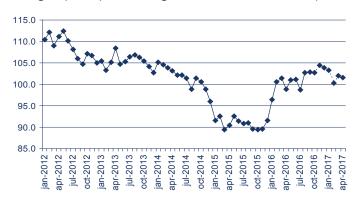


Fig. 45. Employment (million people)

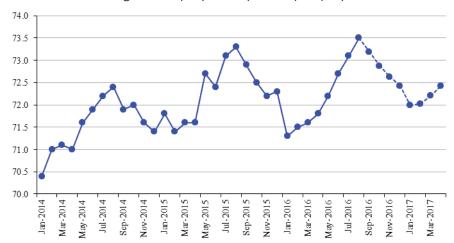
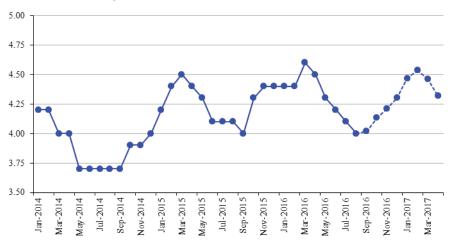


Fig. 46. Unemployment (million people)



Model calculations of short-term forecasts of social and economic indices of the Russian Federation: October 2016

MOUCH CAICHIALIOUS OF SHOTT-UCTIN TOLCCASES OF SOCIAL AND CCONOMIC MIGRICES		I CIIC IVA	of the Nussian Feuel auon.			70107	100	Ŀ	
Index			2016				2017		
	Aug	\mathbf{Sep}	0ct	Nov	Dec	Jan	Feb	Mar	Apr
Rosstat IIIP (growth rate, %)*	0.7	-0.8	-0.3	8.0	-0.3	2.0	0.4	1.1	0.7
HSE IIP (growth rate %)*	1.8	1.6	1.4	2.4	2.1	2.3	1.0	2.0	1.6
Rosstat IIP for mining (growth rate, %)*	1.8	-1.2	-1.3	9.0	-0.2	0.3	-2.2	-1.5	6.0
HSE IIP for mining (growth rate, %)*	1.5	1.8	1.1	1.5	2.2	0.0	-2.7	-2.3	0.3
Rosstat IIIP for manufacturing (growth rate, %)*	0.1	-1.2	6.0-	-0.3	9.0	0.3	0.2	1.8	2.3
HSE IIP for manufacturing (growth rate, %)*	1.9	1.2	1.5	1.8	2.0	0.7	-1.3	0.4	1.3
Rosstat IIP for utilities (electricity, water, and gas) (growth rate, %)*	1.3	3.6	0.3	8.0	4.0	-2.5	8.0	2.5	4.8
HSE for utilities (electricity, water, and gas) (growth rate, %)*	2.4	2.6	8.0	0.7	6.9	-0.9	8.0	6.2	4.0
Rosstat IIP for food products (growth rate, %)*	3.2	2.5	1.3	2.5	1.6	3.4	2.8	2.4	2.5
HSE IIP for food products (growth rate, %)*	3.1	0.2	1.2	1.7	3.0	3.1	1.6	2.0	2.6
Rosstat IIP for coke and petroleum (growth rate, %)*	-2.8	8.0	-0.1	-3.0	-4.0	-1.1	2.1	9.0-	3.1
HSE for coke and petroleum (growth rate, %)*	-3.7	-2.3	1.6	6.0-	-3.3	0.0	-1.8	0.0	-1.2
Rosstat for primary metals and fabricated metal products (growth rate, %)*	-5.5	-6.5	-3.6	-0.4	1.7	-1.1	-2.1	6.0	3.1
HSE IIP for primary metals and fabricated metal products (growth rate, %)*	9.9-	-6.3	-4.4	-3.9	-2.6	1.4	-1.6	0.7	2.3
	1.4	-1.9	-8.8	-1.1	-7.1	9.1	15.3	9.7	20.9
HSE IIP for machinery (growth rate %)*	3.8	3.4	-1.8	-0.2	5.4	18.8	18.7	8.0	0.3
Retail sales, trillion Rb	2.41	2.44	2.48	2.49	3.15	2.15	2.12	2.32	2.29
Real retail sales (growth rate, %)*	-3.6	-3.4	-3.27	-3.2	-2.9	-4.56	-4.95	-4.2	-4.5
Investments in capital assets, trillion Rb	23.3	25.5	24.1	23.6	26.1	15.8	19.0	22.9	22.2
Real investments in capital assets (growth rate, %)*	19.8	22.0	20.1	19.6	22.4	12.7	16.3	20.3	18.1
Export to all countries (billion \$)	18.4	18.2	15.9	16.7	16.3	8.5	14.3	15.5	15.1
Export to countries outside the CIS (billion \$)	16.5	16.3	14.2	15.0	15.0	6.7	11.3	13.8	13.4
Import from all countries (billion \$)	0.0	0.2	0.4	0.5	0.5	1.2	9.0	9.0	0.7
Import from countries outside the CIS (billion \$)	-1.4	-0.1	0.3	0.0	-0.2	-0.1	0.5	1.3	0.7
CPI (growth rate, %)**	8.9-	-3.1	2.2	-4.4	-1.6	-1.5	-8.5	2.0	-3.7
PPI for industrial goods (growth rate, %)**	-0.3	0.7	1.0	0.4	0.2	1.0	0.5	0.3	0.5
PPI for mining (growth rate, %)**	2.9	0.1	9.0	0.3	0.1	6.0	2.7	6.0	-0.1
PPI for manufacturing (growth rate, %)**	9.0	0.3	0.3	0.4	0.7	0.7	1.1	1.2	1.1
PPI for utilities (electricity, water, and gas) (growth rate, %)**	0.3	0.3	0.4	9.0	0.2	0.7	9.0	9.0	0.7
PPI for food products (growth rate, %)**	0.3	0.4	0.5	0.4	0.3	9.0	0.5	8.0	0.7
PPI for the textile and sewing industry (growth rate, %)**	-0.4	0.2	0.1	-0.1	0.5	9.0	9.0	0.5	0.5
PPI for wood products (growth rate, %)**	-3.0	0.0	1.8	1.2	1.5	1.4	6.0	1.2	1.5
PPI for the pulp and paper industry (growth rate, %)**	6.0-	-0.5	-0.3	-0.2	-0.1	-0.1	0.0	0.1	0.2
PPI for coke and petroleum (growth rate, %)**	9.0-	0.2	0.7	6.0	8.0	-0.2	6.0-	0.2	0.4
PPI for the chemical industry (growth rate, %)**	-0.1	0.5	0.3	8.0	6.0	1.2	1.5	9.0	9.0

			2016				2017	17	
Index	Aug	Sep	Oct	Nov	Dec	Jan	Feb	Mar	Apr
PPI for primary metals and fabricated metal products (growth rate, %)**	1.5	0.0	1.4	0.3	6.0	0.7	0.2	0.7	0.7
PPI for machinery (growth rate, %)**	3.62	3.62	3.61	3.62	3.65	3.69	3.72	3.75	3.78
PPI for transport equipment manufacturing (growth rate, %)**	0.4	0.1	0.1	0.0	0.0	1.7	0.0	-0.1	-0.1
The cost of the monthly per capita minimum food basket (thousand Rb)	0.0	2.4	-0.5	9.0	2.3	6.0	-0.2	1.5	1.8
The composite index of transportation tariffs (growth rate, %)**	0.1	0.1	0.1	0.1	0.1	0.1	0.1	0.0	3.8
The index of pipeline tariffs (growth rate, %)**	46.1	46.2	45.3	45.7	44.9	45.2	45.0	44.9	45.2
The index of motor freight tariffs (growth rate, %)**	1.64	1.59	1.60	1.58	1.55	1.55	1.54	1.53	1.53
The Brent oil price (\$ a barrel)	1.34	1.33	1.33	1.35	1.37	1.37	1.37	1.39	1.41
The aluminum price (thousand \$ a ton)	4.75	4.72	4.73	4.74	4.74	4.76	4.77	4.79	4.80
The gold price (thousand \$ per ounce)	10.3	10.2	10.4	10.5	10.5	10.6	10.6	10.6	10.7
The nickel price (thousand \$ a ton)	8.56	8.52	8.56	8.71	89.8	9.18	8.80	8.96	0.89
The copper price (thousand \$ a ton)	36.7	36.8	37.0	37.2	37.5	38.5	38.7	38.9	39.0
The monetary base (trillion Rb)	0.39	0.40	0.40	0.40	0.40	0.40	0.40	0.41	0.41
M2 (trillion Rb)	64.91	63.16	63.00	63.08	63.63	63.84	64.16	64.44	64.70
Gold and foreign exchange reserves (billion \$)	1.11	1.12	1.10	1.11	1.10	1.11	1.11	1.11	1.11
The RUR/USD exchange rate (rubles per one USD)	-8.2	-2.8	-4.5	-2.6	-3.6	-0.8	-2.0	-2.4	0.0
The USD/EUR exchange rate (USD per one Euro)	-7.3	-2.8	-4.1	-2.4	-3.7	-1.3	-1.7	-2.0	-0.1
Real disposable cash income (growth rate, %)*	2.7	2.8	2.8	4.4	3.8	3.3	0.3	2.0	1.6
Real cash income (growth rate, %)*	73.5	73.2	72.9	72.6	72.4	72.0	72.0	72.2	72.4
Real accrued wages (growth rate, %)*	4.0	4.0	4.1	4.2	4.3	4.4	4.5	4.5	4.4
Employment (million people)	72.3	72.0	72.1	72.3	72.4	72.9	72.7	73.0	73.2
Unemployment (million people)	4.4	4.6	3.1	4.5	4.4	4.2	4.2	4.3	9.1

Note. Actual values are printed in the bold type *% of the respective month of the previous year **% of the previous month